Multi-dimensional Approximate Counting*

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Abstract

The celebrated Morris counter uses $\log_2\log_2 n + O(\log_2\sigma^{-1})$ bits to count up to n with a relative error σ , where if $\hat{\lambda}$ is the estimate of the current count λ , then $\mathbb{E}|\hat{\lambda}-\lambda|^2<\sigma^2\lambda^2$. The Morris counter was proved to be optimal in space complexity by Nelson and Yu [14], even when considering the error tails. A natural generalization is multi-dimensional approximate counting. Let $d\geq 1$ be the dimension. The count vector $x\in\mathbb{N}^d$ is incremented entry-wisely over a stream of coordinates $(w_1,\ldots,w_n)\in[d]^n$, where upon receiving $w_k\in[d]$, $x_{w_k}\leftarrow x_{w_k}+1$. A d-dimensional approximate counter is required to count d coordinates simultaneously and return an estimate \hat{x} of the count vector x. Aden-Ali, Han, Nelson, and Yu [1] showed that the trivial solution of using d Morris counters that track d coordinates separately is already optimal in space, if each entry only allows error relative to itself, i.e., $\mathbb{E}|\hat{x}_j-x_j|^2<\sigma^2|x_j|^2$ for each $j\in[d]$. However, for another natural error metric—the Euclidean mean squared error $\mathbb{E}|\hat{x}-x|^2$ —we show that using d separate Morris counters is sub-optimal.

In this work, we present a simple and optimal d-dimensional counter with Euclidean relative error σ , i.e., $\mathbb{E}|\hat{x}-x|^2 < \sigma^2|x|^2$ where $|x| = \sqrt{\sum_{j=1}^d x_j^2}$, with a matching lower bound. We prove the following.

- There exists a $(\log_2 \log_2 n + O(d \log_2 \sigma^{-1}))$ -bit d-dimensional counter with relative error σ .
- Any d-dimensional counter with relative error σ takes at least $(\log_2\log_2n + \Omega(d\log_2\sigma^{-1}))$ bits of space. The upper and lower bounds are proved with ideas that are strikingly simple. The upper bound is constructed with a certain variable-length integer encoding and the lower bound is derived from a straightforward volumetric estimation of sphere covering.

1 Introduction

In 1978, Morris [13] invented the classic approximate counter which can count up to n with $\log_2 \log_2 n + O(\log_2 \sigma^{-1})$ bits, returning an estimate of the current count with a relative error σ . Such approximate counters are invented mainly to save space. However, it was noticed quite recently that approximate counters can also be much faster on modern hardware in comparison to the deterministic counter due to their low write complexity [18, 17, 9].

In the original paper [13], Morris analyzed the mean and variance of the estimates. Flajolet [6] analyzed the mean and variance of the $index^1$ of the Morris counter, from which a quantified space bound can be derived with Chebyshev's bound. Nelson and Yu [14] analyzed the tail of the index, obtaining a sharper dependence on the failure probability. Specifically, for increments up to λ and $\epsilon, \delta > 0$, if an estimate $\hat{\lambda}$ is produced with $\mathbb{P}(|\hat{\lambda} - \lambda| < \epsilon \lambda) \ge 1 - \delta$ then the Morris counter needs $O(\log \log n + \log \epsilon^{-1} + \log \log \delta^{-1})$ space, and this is proved to be optimal [14].²

It is the common scenario in real-world applications that many different counters are maintained simultaneously. For example, the original motivation of Morris is to count the number of each trigram in texts [12, 14] where $d = 26^3$ different trigrams are counted simultaneously. Aden-Ali, Han, Nelson, and Yu [1] showed that, essentially, d independent counters cannot be compressed as long as each counter is required to report an estimate with an error relative to itself. However, as we will later show, the space can be compressed if the error metric is the d-dimensional Euclidean norm. While the per-entry relative error requirement considered in [1] is reasonable when an approximation for each individual entry is needed, we remark that in applications where an approximation of the whole count vector is needed, the Euclidean error metric is more natural.

^{*}This work was supported by NSF Grant CCF-2221980.

¹Roughly speaking, Morris counter stores an index v to represent a count around 2^v . Therefore, to count up to n, the index increases to $\log n$ and thus it takes $\log \log n$ to store the index.

²Technically, one needs to keep a separate deterministic counter for small $\lambda = O(a)$ to obtain the optimal failure rate [14].

We now formally define the problem of d-dimensional counting in terms of Euclidean mean squared error. Let (e_1, e_2, \ldots, e_d) be the standard basis of \mathbb{R}^d where e_k has its kth component being one and other components being zero.

DEFINITION 1.1. ((EUCLIDEAN) d-DIMENSIONAL COUNTING) Let $n \ge 1$ be the maximum stream length, $d \ge 1$ be the dimension, and $\sigma > 0$ be the relative error. An (n,d,σ) -counter is a randomized data structure that, for any input sequence $(w_1,\ldots,w_k)\in [d]^k$ with $k\le n$, returns an estimate $\hat x$ of $x=\sum_{j=1}^k e_{w_j}$ with $\mathbb{E}|\hat x-x|^2<\sigma^2|x|^2$, where $|x|=\sqrt{\sum_{j=1}^d x_j^2}$ is the Euclidean length. We denote the state space of the counter as Ω with $|\Omega|\ge 1$. One counter thus uses $\log_2 |\Omega|$ bits of space.

Recall that the original Morris counter [13] with parameter $a \ge 1$, denoted as Morris(a), returns an unbiased estimate of the count with relative variance O(1/a). Thus by definition, Morris(a) is an $(n, 1, O(\sqrt{1/a}))$ -counter. The error here is measured by the mean squared error $\mathbb{E}|\hat{x}-x|^2$, which is equivalent to measure the variance if the estimator is unbiased. We note that there are, of course, many other reasonable error metrics to consider, depending on the specific applications. The mean squared error is a good starting point for a new algorithmic problem due to its simple probabilistic/geometric structures.

One natural way to construct a d-dimensional counter is to simply use d separate Morris counters, which is the original method that Morris used to count trigrams [12]. Suppose now we have d $(n, 1, \sigma)$ -counters to count x_1, \ldots, x_d . Let the estimates be $\widehat{x_1}, \ldots, \widehat{x_d}$ respectively. Then we have

$$\mathbb{E}|\hat{x} - x|^2 = \mathbb{E}\sum_{j=1}^{d} |\hat{x_j} - x_j|^2 = \sum_{j=1}^{d} \mathbb{E}|\hat{x_j} - x_j|^2$$

Since for any j, $\hat{x_j}$ is an $(n, 1, \sigma)$ -counter for x_j , we have $\mathbb{E}|\hat{x_j} - x_j|^2 < \sigma^2 x_i^2$. Thus,

$$\mathbb{E}|\hat{x} - x|^2 < \sum_{i=1}^d \sigma^2 x_j^2 = \sigma^2 |x|^2.$$

We thus see d $(n,1,\sigma)$ -counters do form an (n,d,σ) -counter (Definition 1.1). Each $(n,1,\sigma)$ -counter can be implemented with a $\operatorname{Morris}(O(\sigma^{-2}))$ counter, taking $\log_2\log_2 n + O(\log_2\sigma^{-1})$ bits. Thus the total space needed is $d\log_2\log_2 n + O(d\log_2\sigma^{-1})$ bits, which is sub-optimal for this task.

The main contribution of this work is the construction of a simple and optimal (n, d, σ) -counter, with a matching lower bound.

Theorem 1.1. The following statements are true.

Upper bound For any $n \geq 2$, $d \geq 1$, and $\sigma \in (0, 1/3)$, there exists an (n, d, σ) -counter with space size $\log_2 |\Omega| = \log_2 \log_2 n + O(d \log_2 \sigma^{-1})$.

Lower bound For any $n \ge e^2$, $d \ge 1$, and $\sigma \in (0, 1/16)$, if there is an (n, d, σ) -counter with state space Ω , then $\log_2 |\Omega| \ge \log_2 \log_2 n + \Omega(d \log_2 \sigma^{-1})$.

A related data structure is the Count-Min sketch [3] by Cormode and Muthukrishnan. Given a count vector $x \in \mathbb{N}^d$ with $\sum_{j=1}^d x_j = n$, Count-Min is able to return an estimate $\widehat{x_j}$ of x_j , for any $j \in [d]$, such that $\widehat{x_j} \in [x_j, x_j + \epsilon n]$ with probability $1 - \delta$, using only $O(\epsilon^{-1} \log \delta^{-1} \log n)$ bits of space. It may seem that Count-Min with the power of hash functions has a chance to beat the lower bound in Theorem 1.1 by combining the estimates for each coordinate. However, it is hopeless to use Count-Min as a d-dimensional counter unless $\epsilon = O(1/d)$, in which case it still lies above the lower bound in Theorem 1.1. Indeed, suppose $\widehat{x_j} \in [x_j, x_j + \epsilon n]$ holds for all j. The vector estimate $\widehat{x} = (\widehat{x_1}, \dots, \widehat{x_d})$ has error

$$|\hat{x} - x|^2 \le d\epsilon^2 n^2 \le d^2 \epsilon^2 |x|^2,$$

where we use the bound $|x|^2 \ge (\sum_{j=1}^d x_j)^2/d = n^2/d$. Therefore ϵ has to be O(1/d) for \hat{x} to be a multiplicative estimate of x.

1.1 Technical Overview

- **1.1.1** Upper Bound The upper bound in Theorem 1.1 uses the natural idea of maintaining a common scale counter U for all d dimensions and tracking the relative magnitude of each coordinate with a vector $V \in \mathbb{N}^d$. The estimate \hat{x} is equal to 2^UV . For demonstration, we will first describe a very simple but sub-optimal solution and then discuss how to modify it to reach optimal space-accuracy tradeoff. The transition of the states is designed similarly with the design of the Morris counter, in which the estimate is maintained unbiased at any moment $(\mathbb{E}2^UV = x)$. Initially we set U = 0 and $V = (0, \dots, 0)$.
 - When the jth coordinate is incremented (i.e., $x_j \leftarrow x_j + 1$), we update $V_j \leftarrow V_j + 1$ with probability 2^{-U} , ensuring the unbiasedness of the estimate.
 - Each coordinate V_j is restricted to the range [0, a], so that the space to store V is $d \log_2(1 + a)$. Whenever some V_j exceeds a, we have to scale up to keep V_j in range. Again, we want to maintain the property that $\mathbb{E}^2 V = x$. Thus, the scale-up should be done as follows.
 - $-U \leftarrow U + 1$. The scale counter is increased by 1.
 - For the estimate to be unbiased, V_j should become $V_j/2$ for every j. Here we also want V_j to always be an integer. Therefore, if V_j is even, then $V_j \leftarrow V_j/2$. If V_j is odd, then $V_j \leftarrow (V_j + \xi)/2$ where $\xi \in \{-1,1\}$ is a freshly sampled Rademacher random variable.

The algorithm above is unbiased by design. However, the memory-accuracy tradeoff is *not* optimal. There is a fundamental problem of the design above that the parameter a has to be $\Omega(\sqrt{d})$ (with $\sigma = O(1)$) for it to be a d-dimensional counter. Instead of analyzing the algorithm in details, this problem can be identified by simply checking the set of *all possible estimates* that are produced by the algorithm:

$$\mathcal{D} = \{2^{U}(V_1, \dots, V_d) : U \in \mathbb{N}, \forall j \in [d], V_j \in [0, a]\}.$$

It turns out that \mathcal{D} is not dense enough for estimating some input x. Indeed, we choose

$$x = 2^{s}(\underbrace{a, \dots, a}_{r}, \underbrace{1/2, \dots, 1/2}_{d-r}),$$

for some $r \in [0,d]$ so that both $y_1 = 2^s(a,\dots,a,1,\dots,1)$ and $y_2 = 2^{s-1}(a,\dots,a,1,\dots,1)$ are far from x. By construction, the nearest estimate of x in \mathcal{D} has to be either y_1 or y_2 .³ We choose r so that the two choices have equal distances to x. A simple calculation shows that we need to set $r = d/(1+4a^2)$.⁴ Assume now $a < \sqrt{d}/3$ and then r > 1. Thus $\min_{y \in \mathcal{D}} |x - y|^2 = 2^{2s} \frac{da^2}{1+4a^2}$, while $|x|^2 = 2^{2s} \cdot 2a^2d/(1+4a^2)$. Therefore, if the algorithm is an (n,d,σ) -counter, then $\mathbb{E}|\hat{x}-x|^2 < \sigma^2|x|^2$, which means there is at least one $x_* \in \mathcal{D}$ such that $|x_*-x|^2 < \sigma^2|x|^2$. This suggests that $2^{2s}a^2d/(1+4a^2) < \sigma^22^{2s} \cdot 2a^2d/(1+4a^2)$ and therefore $\sigma > 1/\sqrt{2}$. In other words, if $a < \sqrt{d}/3$, then any (n,d,σ) -counter must have $\sigma > 1/\sqrt{2}$. We conclude that a needs to be $\Omega(\sqrt{d})$ when $\sigma \le 1/\sqrt{2}$, indicating an additional $O(d\log d)$ space overhead in comparison to the optimal space.

From the discussion above, we see that one has to have a dense enough set of estimates \mathcal{D} for a correct (n, d, σ) -counter. Roughly speaking, one wants to spend more bits on large coordinates and fewer bits on small coordinates. This can be done with a simple and natural algorithmic idea: variable-length integer encoding. The usual binary encoding of non-negative integers automatically uses more bits on large numbers and fewer bits on small numbers. We will prove the following set of estimates is dense enough.

$$\mathcal{D}_{\text{compressed}} = \{2^U(V_1, \dots, V_d) : U \in \mathbb{N}, \text{the encoding length of } V \text{ is at most } O(d \log_2 \sigma^{-1})\},$$

This leads to our upper bound in Theorem 1.1. One may observe how this encoding idea better handles the previous counterexample $x = 2^s(a, \ldots, a, 1/2, \ldots, 1/2)$. An estimate of $2^{s-1}(2a, \ldots, 2a, 1, \ldots, 1)$ can now be returned because though each entry with value 2a needs above average space to encode, each entry with value 1 uses less than average space to encode, saving the memory space for the large entries.

³Note that y_1 is one possible probabilistic rounding of x. Other roundings $\{2^s(a,\ldots,a,z_1,\ldots,z_{d-r})\mid z_j\in\{0,1\}\}$ all have the same distance to x. By symmetry, we only need to consider y_1 . The vector y_2 is produced by decreasing the scale to 2^{s-1} , so the smaller entries can be represented precisely with the cost of underestimating the large entries.

⁴For simplicity, we only require r > 1 and omit the details of rounding r to an integer.

1.1.2 Lower Bound As we have discussed above, the set of all estimates produced by the counter has to be dense enough for the task of *d*-dimensional counting to be possible. The lower bound in Theorem 1.1 thus arises from estimating the size of *multiplicative space coverings*.

DEFINITION 1.2. (MULTIPLICATIVE SPACE COVERING) For any subsets $A, R \subset \mathbb{R}^d$ and $\sigma > 0$, we say R is a σ -multiplicative covering of A if for any $x \in A$ there exists $y \in R$ such that $|x - y| < \sigma |x|$.

Clearly, the set of all estimates of an (n, d, σ) -counter will form a σ -multiplicative space covering of $\mathcal{X} = \{x \in \mathbb{N}^d \mid x_1 + \dots + x_d \leq n\}$. The lower bound is thus proved by estimating the covering size of \mathcal{X} , which can be further reduced to the classic problem of covering spheres with smaller spheres [16].

1.2 Related Work There have been many variants and analysis of Morris counters before [10, 11, 8, 7, 14]. Nevertheless, they have only considered the one-dimensional case. Tracking a multi-dimensional count vector (with a vector norm) is a new problem to the best of our knowledge.

In the one-dimensional case (d=1), our algorithm is equivalent to the variant—floating-point counter—by Csűrös [4], except for that the floating-point counter restricts a to be a power of two so no probabilistic rounding is needed when halving. Nevertheless, when $d \geq 2$, probabilistic rounding is inevitable. In comparison to the classic Morris counter, the floating-point variant is much more convenient to implement on binary machines, where only simple integer operations and bit operations are needed [4]. Another advantage is that the number of random bits needed per increment is at most 2 in expectation, since the probability 2^{-U} can be simulated by generating random bits sequentially, looking for U consecutive ones [4] (stop generating random bits when a zero shows up). Such advantages are inherited by our d-dimensional counter as well.

1.3 Organization We will construct the upper bound in §2 and prove the lower bound in §3. The main theorem (Theorem 1.1) follows directly from Corollary 2.1 and 3.2. We have included a sample run in Appendix A to help illustrate the patterns of the new algorithm.

2 Upper Bound: Variable-length Integer Encoding

We consider a specific integer encoding which is carefully designed to simplify the analysis later.

DEFINITION 2.1. (VARIABLE-LENGTH INTEGER ENCODING) Let the symbol set be $\{0,1,|\}$ where | is used as a separator. We encode 0 as |, 1 as 0|. For $k \geq 2$, we encode k as $[k-1]_2|$, where $[k-1]_2$ is the binary representation of k-1. See the following table.

integer	code	integer	code
0		4	11
1	0	5	100
2	1	6	101
3	10	7	110

For any $k \in \mathbb{N}$, we define $\psi(k)$ to be the length of its encoding above. An integer vector $V \in \mathbb{N}^d$ is encoded by concatenating the codes for each of its component. For example V = (3,0,4,0,1) is encoded as 10||11||0|. The code length of the vector V is denoted by $\psi(V) = \sum_{j=1}^{d} \psi(V_j)$.

REMARK 2.1. For clarity, the code here uses an alphabet of size 3, including a separator. If it is encoded with bits, then the length will increase by a $\log_2 3$ factor. See [15, 5] for techniques that encode m_* {0,1,|}-symbols using $\lceil m_* \log_2 3 \rceil$ bits with constant querying time for each entry.

We now formally describe the new algorithm. The algorithm stores a scale counter $U \in \mathbb{N}$ and a relative vector $V = (V_1, \dots, V_d) \in \mathbb{N}^d$. The estimator is $\hat{x} = 2^U V = (2^U V_1, \dots, 2^U V_d)$. As discussed in the introduction, we will use the variable-length encoding in Definition 2.1 to store the vector V. Let M_* be the memory budget for V. Recall that $\psi(V)$ is the encoding length of V.

Initialization $(V_1, \ldots, V_d) \leftarrow (0, \ldots, 0)$ and $U \leftarrow 0$.

Increment(j) With probability 2^{-U} , $V_i \leftarrow V_i + 1$. If $\psi(V) > M_*$, then execute Scale-up.

Scale-up $U \leftarrow U + 1$ and for $k \in \{1, \dots, d\}$,

- if V_k is even, then $V_k \leftarrow V_k/2$;
- if V_k is odd, then $V_k \leftarrow (V_k + \xi)/2$ where $\xi \in \{-1,1\}$ is a freshly sampled Rademacher random variable.

Query Return $2^U(V_1,\ldots,V_d)$.

Since the algorithm will execute Scale-up whenever the space $\psi(V)$ of storing V exceeds the budget M_* , the behavior of the algorithm will depend on the way that V is encoded. The codes in Definition 2.1 are constructed so that the space $\psi(V)$ increases by at most one during Increment and decreases by at least one during Scale-up. This guarantees that the memory space to encode V is always bounded by M_* . We now list and prove all the properties of the encoding in Definition 2.1 that we will use in the analysis.

LEMMA 2.1. For any $V \in \mathbb{N}^d$, the following statements are true.

- $\psi(V) \le 2d + \sum_{j=1}^{d} \log_2(1+V_j)$. (Space bound.)
- For any $j \in [d]$, $\psi(V + e_i) \leq \psi(V) + 1$. (An increment increases the space by at most 1.)
- $\psi(|V/2|) \ge \psi(V) 2d$. (Halving the relative vector will decrease the space by at most 2d.)
- If there exists $j \in [d]$ such that $V_i \geq 3$, then $\psi(\lceil V/2 \rceil) \leq \psi(V) 1$. (The space will decrease by at least one if not all V_i s are small.)

Remark 2.2. |V/2| and |V/2| are evaluated entry-wisely.

Proof. Recall that for the usual binary encoding, an integer k has length $1 + |\log_2 k|$ if $k \ge 1$. Thus, by Definition 2.1, if $k \geq 2$, then $2 + \lfloor \log_2(k-1) \rfloor$ symbols are needed to store the code $\lfloor k-1 \rfloor_2 \rfloor$. It requires one symbol to encode 0 and two symbols to encode 1. We bound the length for encoding k by $2 + \log_2(1 + k)$, which holds for all $k \in \mathbb{N}$. We then have $\psi(V) = \sum_{j=1}^{d} \psi(V_j) \leq 2d + \sum_{j=1}^{d} \log_2(1+V_j)$. The second statement is true by construction.

For the third statement, it suffices to prove that $\psi(\lfloor k/2 \rfloor) \geq \psi(k) - 2$ for any $k \in \mathbb{N}$. One may check that this is true for k = 1, 2, 3, 4. Now assume $k \ge 5$. Note that $\psi(|k/2|) - \psi(k) = |\log_2(|k/2| - 1)| - |\log_2(k-1)| =$ $-2+\lfloor \log_2(4\lfloor k/2 \rfloor -4) \rfloor - \lfloor \log_2(k-1) \rfloor$. It suffices to prove that $4\lfloor k/2 \rfloor -4 \ge k-1$. Indeed, $4\lfloor k/2 \rfloor -4 \ge 2k-6 \ge k-1$ since k > 5.

For the fourth statement, since $||V||_{\infty} \geq 3$, there exists $j \in [d]$, such that $V_j \geq 3$. If V_j is even, then $V_j \geq 4$ and $\psi(V_j/2) = 2 + \lfloor \log_2(V_j/2 - 1) \rfloor = 1 + \lfloor \log_2(V_j - 2) \rfloor \leq \psi(V_j) - 1$. If V_j is odd, then $\psi(\lceil V_j/2 \rceil) = \psi((V_j+1)/2) = 2 + \lfloor \log_2((V_j+1)/2-1) \rfloor = 1 + \lfloor \log_2((V_j-1)) \rfloor = \psi(V_j) - 1$. Note that ψ is non-decreasing and thus we conclude that $\psi(\lceil V/2 \rceil) \leq \psi(V) - 1$.

Next, we analyze the variance of the algorithm. Similar to the analysis of the Morris counter [13], the basic idea is to analyze the change in variance caused by each Increment, including a possible Scale-up step, and then compute the final variance by summing up the changes at time k = 1, ..., n. Recall that for any $r \in [d]$, e_r is the rth standard base vector of \mathbb{R}^d .

Theorem 2.1. Fix any input sequence $(w_1, \ldots, w_n) \in [d]^n$. Let (U, V) be the final state of the algorithm and $x = \sum_{j=1}^{n} e_{w_j}$ be the final count vector. Define $\hat{x} = 2^U V$. If the memory budget is $M_* = 4d + 2d \log_2(1+a)$ with $a > \overline{1}$, then

$$\mathbb{E}\hat{x} = x,$$

$$\mathbb{E}|\hat{x} - x|^2 \le \frac{5}{6a - 2}|x|^2.$$

Furthermore for any r > 1,

$$\mathbb{P}\left(U \ge r + \log_2\left(\frac{n}{ad} + 1\right)\right) \le 2^{-r}.$$

Proof. $\mathbb{E}\hat{x} = x$ is true by construction, since both Increment and Scale-up are designed to keep $\hat{x} = 2^U V$ unbiased. The challenge is to analyze the variance $\mathbb{E}|\hat{x} - x|^2$. We start by proving the following lemma that bounds the final variance by the distributional properties of the scale counter U at time $k = 1, \ldots, n$.

LEMMA 2.2. (VARIANCE DECOMPOSITION) For any $k \leq n$, let $(U^{(k)}, V^{(k)})$ be the memory state at time k (i.e., after seeing w_1, \ldots, w_k) and $x^{(k)} = \sum_{j=1}^k e_{w_j}$ be the count vector at time k. Let $\hat{x}^{(k)} = 2^{U^{(k)}}V^{(k)}$ to be the estimate at time k. Then

$$\mathbb{E}|\hat{x}^{(n)}|^2 - |x^{(n)}|^2 \le \sum_{k=0}^{n-1} \mathbb{E}(2^{U^{(k)}} - 1) + \frac{d}{3}\mathbb{E}(2^{2U^{(n)}} - 1).$$

Proof. We define $Y^{(k)}$ to be the intermediate vector after Increment before Scale-up at time k. By construction, we know $Y^{(k)} = V^{(k)} + e_{w_{k+1}}$ with probability $2^{-U^{(k)}}$; otherwise $Y^{(k)} = V^{(k)}$. For simplicity, we now fix $k \ge 0$ and write $U^{(k)}, V^{(k)}, x^{(k)}, w_{k+1}, Y^{(k)}$ as U, V, x, w, Y and $U^{(k+1)}, V^{(k+1)}, x^{(k+1)}$ as U', V', x' respectively. By the definition of Increment, we have

(2.1)
$$\mathbb{E}(|2^{U}Y|^{2} - |2^{U}V|^{2} | U, V) = 2^{-U} \cdot 2^{2U}(2V_{w} + 1) = 2^{U}(2V_{w} + 1).$$

Now if $\psi(Y) \leq M_*$, then we have (U',V') = (U,Y) where no Scale-up is executed. If $\psi(Y) > M_*$, then we must have $\psi(Y) = M_* + 1$ because by Lemma 2.1, $\psi(Y) \leq \psi(V + e_w) \leq \psi(V) + 1$ and $\psi(V) \leq M_*$. By construction, $M_* \geq 3d$ and thus there exists $j \in [d]$ such that $Y_j \geq 3$. By Lemma 2.1, we know $\psi(\lceil Y/2 \rceil) \leq \psi(Y) - 1 = M_*$. Therefore, after the Scale-up step, the space will be at most M_* . When Scale-up is executed, U' = U + 1 and $V' = Y/2 + \sum_{k=1}^d \xi_k \mathbbm{1}[2 \nmid Y_k]$. Here, ξ_1, \ldots, ξ_d are i.i.d. Rademacher random variables. Thus we have

$$\begin{split} & \mathbb{E}(|2^{U'}V'|^2 - |2^UY|^2 \mid U, Y) \\ & = \mathbb{1}\left[\psi(Y) > M_*\right] \sum_{k=1}^d \mathbb{E}\left[\left(2^{2U+2} \left(\frac{Y_k + \xi_k \mathbb{1}\left[2 \nmid Y_k\right]}{2}\right)^2 - 2^{2U}Y_k^2\right) \mid U, Y\right], \end{split}$$

note that $\mathbb{E}\xi_k = 0$ and $\xi_k^2 = 1$

$$= \mathbb{1} \left[\psi(Y) > M_* \right] \sum_{k=1}^d \left(2^{2U} (Y_k^2 + \mathbb{1} \left[2 \nmid Y_k \right]) - 2^{2U} Y_k^2 \right)$$

$$= \mathbb{1} \left[\psi(Y) > M_* \right] 2^{2U} \sum_{k=1}^d \mathbb{1} \left[2 \nmid Y_k \right]$$

$$\leq \mathbb{1} \left[\psi(Y) > M_* \right] 2^{2U} \cdot d.$$
(2.2)

Take the total expectation and by Equation (2.1) and (2.2) we have,

$$\mathbb{E}(|2^{U'}V'|^2 - |2^{U}V|^2) = \mathbb{E}\mathbb{E}(|2^{U}Y|^2 - |2^{U}V|^2 \mid U, V) + \mathbb{E}\mathbb{E}(|2^{U'}V'|^2 - |2^{U}Y|^2 \mid U, Y)$$

$$\leq \mathbb{E}(2^{U}(2V_w + 1)) + \mathbb{E}(\mathbb{1} \left[\psi(Y) > M_*\right] 2^{2U} d).$$
(2.3)

Recall that $|x'|^2 - |x|^2 = (x_w + 1)^2 - x_w^2 = 2x_w + 1$ and $\mathbb{E}2^U V_w = x_w$. This implies $\mathbb{E}2^U V_w = (|x'|^2 - |x|^2 - 1)/2$. Thus by Equation (2.3) we have,

$$(2.4) \mathbb{E}(|2^{U'}V'|^2 - |2^{U}V|^2) \le |x'|^2 - |x|^2 + \mathbb{E}(2^{U} - 1) + \mathbb{E}(d2^{2U}\mathbb{1}[\psi(Y) > M_*]).$$

Sum up the differences and we have

$$\mathbb{E}|2^{U^{(n)}}V^{(n)}|^2 = \sum_{k=0}^{n-1} \mathbb{E}(|2^{U^{(k+1)}}V^{(k+1)}|^2 - |2^{U^{(k)}}V^{(k)}|^2)$$

Recall that U', V', U, V are short for $U^{(k+1)}, V^{(k+1)}, U^{(k)}, V^{(k)}$. By Equation (2.4),

$$\leq \sum_{k=0}^{n-1} \left(|x^{(k+1)}|^2 - |x^{(k)}|^2 + \mathbb{E}(2^{U^{(k)}} - 1) + \mathbb{E}\left(d2^{2U^{(k)}} \mathbb{1}\left[\psi(Y^{(k)}) > M_* \right] \right) \right)$$

$$= |x^{(n)}|^2 + \sum_{k=0}^{n-1} \mathbb{E}(2^{U^{(k)}} - 1) + d\mathbb{E}\left[\sum_{k=0}^{n-1} 2^{2U^{(k)}} \mathbb{1}\left[\psi(Y^{(k)}) > M_* \right] \right].$$

Note that $\mathbb{1}\left[\psi(Y^{(k)})>M_*\right]$ is the indicator whether Scale-up is executed when processing w_{k+1} . Therefore

$$\sum_{k=0}^{n-1} 2^{2U^{(k)}} \mathbbm{1} \left[\psi(Y^{(k)}) > M_* \right] = \sum_{j=0}^{U^{(n)}-1} 2^{2j} = \frac{4^{U^{(n)}}-1}{4-1} = \frac{2^{2U^{(n)}}-1}{3}$$

We conclude that

$$\mathbb{E}|2^{U^{(n)}}V^{(n)}|^2 - |x^{(n)}|^2 \le \sum_{k=0}^{n-1} \mathbb{E}(2^{U^{(k)}} - 1) + \frac{d}{3}\mathbb{E}(2^{2U^{(n)}} - 1),$$

which gives the stated variance decomposition.

The next step is to estimate $\mathbb{E}(2^U - 1)$ and $\mathbb{E}(2^{2U} - 1)$. The idea is to relate them with $\mathbb{E}2^U V = x$ and $\mathbb{E}|2^U V|^2 = \mathbb{E}|\hat{x}|^2$.

LEMMA 2.3. If $U \ge 1$, then $|V|^2 \ge \sum_{j=1}^d V_j \ge ad$.

Proof. Since V_j s are integer, we automatically have $|V|^2 = \sum_{j=1}^d V_j^2 \ge \sum_{j=1}^d V_j$. Thus it suffices to prove $\sum_{j=1}^d V_j \ge ad$. Note that if $U \ge 1$, then the algorithm has executed at least one Scale-up. Note that by Lemma 2.1, the space $\psi(V)$ increases at most by one per Increment and decreases at most by 2d per Scale-up. Therefore, after the first Scale-up, the memory usage $\psi(V)$ is between $M_* + 1 - 2d$ and M_* . Note that M_* is set to $4d + d\log_2(1+a)$. Thus we have

$$\psi(V) \ge M_* + 1 - 2d \ge 2d + d\log_2(1+a).$$

On the other hand, by Lemma 2.1 (encoding space bound), we have

$$\psi(V) \le 2d + \sum_{j=1}^{d} \log_2(1 + V_j).$$

The bounds above imply that $\sum_{j=1}^{d} \log_2(1+V_j) \ge d \log_2(1+a)$. Since $\log_2(1+r)$ is convex in r, we conclude that

$$d\log_2(1+a) \le \sum_{j=1}^d \log_2(1+V_j) \le d\log_2\left(1+\sum_{j=1}^d V_j/d\right),$$

and thus $\sum_{j=1}^{d} V_j \ge ad$.

Note that if U=0, then $2^U-1=0$. With the lemma above, we now have

$$ad\mathbb{E}(2^U - 1) \le \mathbb{E}(2^U - 1) \sum_{j=1}^d V_j \le \mathbb{E}2^U \sum_{j=1}^d V_j = \sum_{j=1}^d x_j = k,$$

where k is the current time, i.e., the total number of increments so far. This implies $\mathbb{E}(2^U - 1) \leq k/(ad)$. Similarly, we have

$$ad\mathbb{E}(2^{2U}-1) \leq \mathbb{E}(2^{2U}-1) \sum_{j=1}^d V_j^2 \leq \mathbb{E}2^{2U} \sum_{j=1}^d V_j^2 = \mathbb{E}|\hat{x}|^2,$$

which implies $\mathbb{E}(2^{2U}-1) \leq \mathbb{E}|\hat{x}|^2/(ad)$. Inserting the estimates back into Lemma 2.2, we have

$$\mathbb{E}|\hat{x}^{(n)}|^2 - |x^{(n)}|^2 \le \sum_{k=0}^{n-1} \mathbb{E}(2^{U^{(k)}} - 1) + \frac{d}{3}\mathbb{E}(2^{2U^{(n)}} - 1)$$

$$\le \sum_{k=0}^{n-1} k/(ad) + \frac{d}{3}\mathbb{E}|\hat{x}^{(n)}|^2/(ad)$$

Note that $\sum_{k=0}^{n-1} k \le n^2/2 \le d|x^{(n)}|^2/2$ since $\sum_{j=1}^d x_j^{(n)} = n$ and $d|x^{(n)}|^2 \ge (\sum_{j=1}^d x_j^{(n)})^2$.

$$\leq \frac{1}{2a}|x^{(n)}|^2 + \frac{1}{3a}\mathbb{E}|\hat{x}^{(n)}|^2$$

Reorganizing the inequality above, we have

$$\mathbb{E}|\hat{x}^{(n)}|^2 - |x^{(n)}|^2 \le \frac{5}{6a-2}|x^{(n)}|^2.$$

Note that since the estimate is unbiased, we have $\mathbb{E}|\hat{x}^{(n)}|^2 - |x^{(n)}|^2 = \mathbb{E}|\hat{x}^{(n)} - x^{(n)}|^2$. Thus we conclude that $\mathbb{E}|\hat{x}^{(n)} - x^{(n)}|^2 \le \frac{5}{6a-2}|x^{(n)}|^2$.

Finally we estimate the tail of U. We already know that $\mathbb{E}(2^{U^{(k)}}-1) \leq k/(ad)$ and thus $\mathbb{E}2^{U^{(k)}} \leq 1+k/(ad)$. Then for any z>0, by Markov,

$$\mathbb{P}(2^{U^{(k)}} \ge z) \le \frac{\mathbb{E}2^{U^{(k)}}}{z} \le \left(\frac{k}{ad} + 1\right)/z$$

Setting $z = 2^r(k/(ad) + 1)$ and we have

$$\mathbb{P}\left(U^{(k)} \ge r + \log_2\left(\frac{k}{ad} + 1\right)\right) \le 2^{-r}.$$

We thus have proved the tail bound for U.

Theorem 2.1 characterizes the mean and variance of the estimator. We now analyze the space. The vector V is encoded with length at most M_* by construction. We are left to analyze the memory space needed for the scale counter U.

COROLLARY 2.1. For $\sigma \in (0, 1/3)$, there exists an (n, d, σ) -counter with state space size $\log_2 |\Omega| = \log_2 \log_2 n + O(d \log_2 \sigma^{-1})$.

Proof. If $n \leq \sigma^{-1}$, we may use d deterministic counters taking $d \log_2 n \leq d \log_2 \sigma^{-1}$ space which surely form a (n, d, σ) -counter. We thus assume $n > \sigma^{-1}$.

With a fixed space budget for storing the scale counter U, there is a chance that U grows too large to be stored. In such case, the algorithm simply enters a fail state and returns $(0, \ldots, 0)$ upon query. For $\tau > 0$, we

store U only up to $U_* = \log_2(\tau^{-1}) + \log_2(n/(ad) + 1)$. By design U is non-decreasing in time and it suffices to analyze U at the end of the stream. By Theorem 2.1, we know $\mathbb{P}(U \geq U_*) \leq \tau$. The estimate \hat{x} can be written as $\mathbb{1}[U < U_*] 2^U V$. Its mean squared error is

$$\mathbb{E} \left| \mathbb{1} \left[U < U_* \right] 2^U V - x \right|^2 = \mathbb{E} \left(\mathbb{1} \left[U < U_* \right] |2^U V - x|^2 + \mathbb{1} \left[U \ge U_* \right] |x|^2 \right)$$

$$\leq \mathbb{E} |2^U V - x|^2 + \mathbb{P} (U \ge U_*) |x|^2$$

Note that $\mathbb{E}|2^UV-x|^2 \leq \frac{5}{6a-2}|x|^2$ by Theorem 2.1 and $\mathbb{P}(U \geq U_*) \leq \tau$ by construction.

$$\leq \left(\frac{5}{6a-2} + \tau\right)|x|^2$$

We set $\tau = \sigma^2/2$ and $a = 2\sigma^{-2}$ and have

$$\frac{5}{6a-2} + \tau = \frac{5}{12\sigma^{-2} - 2} + \sigma^2/2 \le \sigma^2,$$

since $\sigma < 1$. In this setting, we have $\mathbb{E}[\mathbb{1}[U < U_*] 2^U V - x|^2 \le \sigma^2 |x|^2$ which implies the algorithm is an (n,d,σ) -counter. We now compute its space usage. By construction, the space used by U is at most $\log_2(\log_2(\tau^{-1}) + \log_2(n/(da) + 1))) \le \log_2\log_2 n + O(1)$ bits since we assumed $\tau^{-1} < n$. The space used by the vector V is by construction $M_* = 4d + d\log_2(1+a) = O(d\log\sigma^{-1})$. Note that V is encoded with a alphabet set of size 3 $(\{0,1,|\})$ and M_* is the maximum number of symbols. One may encode V using bits by adding an additional $\log_2 3$ factor to the leading constant. We conclude that the total space needed to store (U,V) is $\log_2\log_2 n + O(d\log\sigma^{-1})$. \square

3 Lower Bound: Multiplicative Space Covering

Recall that by Definition 1.2, for any subset $A, R \subset \mathbb{R}^d$ and $\sigma > 0$, we say R is an σ -multiplicative covering of A if for any $x \in A$ there exists $y \in R$ such that $|x - y| < \sigma |x|$.

We now prove that any correct (n, d, σ) -counter will induce an $O(\sigma)$ -multiplicative covering of the space. The proof here will be slightly more complicated than what we described in §1.1.2. Previously we assumed the algorithm returns a fixed estimate for a given state and therefore the set of estimates naturally form a covering. In general, a probabilistic algorithm is allowed to produce a randomized answer even if the final memory state is fixed. We show that even when estimates are randomized for a given state, an (n, d, σ) -counter will still induce a covering of the space.

LEMMA 3.1. If Ω is the state space of an (n,d,σ) -counter with $\sigma < 1/3$, then there exists a 4σ -multiplicative covering R of $\{x \in \mathbb{R}^d_+ \mid \sqrt{d}\sigma^{-1} \leq |x| \leq n/\sqrt{d}\}$ with $|R| \leq |\Omega|$.

Proof. Let $g: \Omega \to \mathbb{R}^d$ be the randomized function that returns the estimate $\hat{x} = g(S)$ based on the final state $S \in \Omega$. Let $\mathcal{X} = \{x \in \mathbb{N}^d \mid \sum_{j=1}^d x_j \leq n\}$ be the space of all possible count vectors. By definition of the (n, d, σ) -counter, for any $x \in \mathcal{X}$, we have

$$\mathbb{E}|g(S) - x|^2 < \sigma^2 |x|^2.$$

For any $x \in \mathcal{X}$, define $s_x = \arg\min_{s \in \Omega} \mathbb{E}|g(s) - x|^2$. Thus we have

(3.5)
$$\mathbb{E}|g(s_x) - x|^2 \le \mathbb{E}|g(S) - x|^2 < \sigma^2|x|^2.$$

This implies $\mathbb{E}|g(s_x) - x| \leq \sqrt{\mathbb{E}|g(s_x) - x|^2} < \sigma|x|$. Suppose for some $y \in \mathcal{X}$ with $s_y = s_x$. By using the same randomness for $g(s_x)$ and $g(s_y)$, we have

$$|x - y| = |g(s_x) - x - (g(s_y) - y)| \le |g(s_x) - x| + |(g(s_y) - y)|$$

By Equation (3.5), we know $\mathbb{E}|g(s_x) - x| < \sigma |x|^2$ and $\mathbb{E}|g(s_y) - y| < \sigma |y|^2$. Take expectation on both sides of Equation (3.6) and we have

$$|x - y| \le \mathbb{E}|g(s_x) - x| + \mathbb{E}|(g(s_y) - y)| < \sigma(|x| + |y|).$$

We now group the count vectors x by the value s_x . For any $s \in \Omega$, let $A_s = \{x \in \mathcal{X} \mid s_x = s\}$ and $\{A_s | s \in \Omega\}$ form a partition of \mathcal{X} . For each s that $A_s \neq \emptyset$, we choose a representation $x_*^{(s)} \in A_s$ according to the lexicographic order. Next we are going to show that $R = \{x_*^{(s)} \mid s \in \Omega, A_s \neq \emptyset\}$ forms a good multiplicative covering of $\{x \in \mathbb{R}^d_+ \mid d\sigma^{-1} < |x| < n/d\}$.

For any $x \in \mathbb{R}^d_+$, denote $\lfloor x \rfloor = (\lfloor x_1 \rfloor, \dots, \lfloor x_d \rfloor)$. We have $|x - \lfloor x \rfloor| < \sqrt{d}$. Note that

$$\sum_{j=1}^{d} \lfloor x_j \rfloor \le \sqrt{d} |\lfloor x \rfloor| \le \sqrt{d} |x|.$$

If $|x| \leq n/\sqrt{d}$, then $\sum_{j=1}^{d} \lfloor x_j \rfloor \leq \sqrt{d}|x| \leq n$, which implies that $\lfloor x \rfloor \in \mathcal{X}$. Set $y = x_*^{(s_{\lfloor x \rfloor})} \in R$. Since $s_y = s_{\lfloor x \rfloor}$, we have by Equation (3.7),

$$|y - \lfloor x \rfloor| < \sigma(|y| + |\lfloor x \rfloor|) \le \sigma(|\lfloor x \rfloor| + |y - \lfloor x \rfloor| + |\lfloor x \rfloor|)$$

Reorganizing,

$$|y - \lfloor x \rfloor| < \frac{2\sigma}{1 - \sigma} |\lfloor x \rfloor|.$$

Thus we have

$$|y - x| \le |y - \lfloor x \rfloor| + |x - \lfloor x \rfloor|$$

$$< \frac{2\sigma}{1 - \sigma} |\lfloor x \rfloor| + \sqrt{d}$$

$$\le \frac{2\sigma}{1 - \sigma} |x| + \sqrt{d},$$

If $|x| \geq \sigma^{-1} \sqrt{d}$, then $\sqrt{d} \leq \sigma |x|$. Thus we have

$$|y - x| < \frac{2\sigma}{1 - \sigma}|x| + \sigma|x| = \frac{3\sigma - \sigma^2}{1 - \sigma}|x|.$$

Since $\sigma < 1/3$, we have $\frac{3\sigma - \sigma^2}{1 - \sigma} \le 4\sigma$. Thus $|y - x| < 4\sigma|x|$. Note that the above argument holds for any x with $|x| \in [\sigma^{-1}\sqrt{d}, n/\sqrt{d}]$. Therefore R is a 4σ -multiplicative cover of $\{x \in \mathbb{R}^d_+ \mid \sigma^{-1}\sqrt{d} \le |x| \le n/\sqrt{d}\}$. Finally, note that the covering $R = \{x_*^{(s)} \mid s \in \Omega, A_s \neq \emptyset\}$ has size at most $|\Omega|$.

Lemma 3.1 shows that the state space of an (n, d, σ) -counter will be at least the size of an optimal 4σ -multiplicative covering of the shell⁵ $\{x \in \mathbb{R}^d_+ \mid \sqrt{d}\sigma^{-1} \leq |x| \leq n/\sqrt{d}\}$. We now bound the size of the shell covering. The first step is to bound the size of the sphere covering. Covering a sphere with smaller spheres is a classic problem [16]. We prove the following lemma based on [2].

LEMMA 3.2. For any $\sigma \in (0,1/3)$, if $R \subset \mathbb{R}^d$ is an σ -multiplicative cover of $\{x \in \mathbb{R}^d_+ \mid |x| = 1\}$, then $|R| \geq 2^{-d}\sigma^{-(d-1)}$.

Proof. We first consider the case d=1. The target set $\{x \in \mathbb{R}_+ \mid |x|=1\}=\{1\}$ has size 1 and thus $|R|=1>2^{-1}$. We now assume $d \geq 2$.

Let \mathbb{S}^{d-1} be the unit sphere in \mathbb{R}^d . Let $|\mathbb{S}^{d-1}|$ be the surface area of the sphere. Let the maximum area on \mathbb{S}^{d-1} that can be covered by a single ball with radius σ be η_{σ} . By Corollary 3.2 in [2],

$$\frac{|\mathbb{S}^{d-1}|}{n_{\sigma}} \ge \sqrt{2\pi(d-1)} \frac{\sqrt{1-\sigma^2}}{\sigma^{(d-1)}},$$

if $\sqrt{1-\sigma^2} \ge 1/\sqrt{d}$, which is true since we assumed $\sigma < 1/3$ and $d \ge 2$.

⁵Technically, the shell intersected with the positive cone.

On the other hand, by symmetry, $|\mathbb{S}^{d-1} \cap \mathbb{R}^d_+| = 2^{-d} |\mathbb{S}^{d-1}|$. Therefore, the number of points in R is at least

$$\frac{|\mathbb{S}^{d-1} \cap \mathbb{R}^d_+|}{\eta_{\sigma}} \geq 2^{-d} \sqrt{2\pi (d-1)} \frac{\sqrt{1-\sigma^2}}{\sigma^{(d-1)}} \geq 2^{-d} \sigma^{-(d-1)},$$

since $d \ge 2$ and $\sqrt{2\pi(1-1/9)} \ge 1$.

With the bound for the sphere covering in hand, the next step is to count the number of *layers* in the shell so that the coverings for different layers are disjoint. Define $\mathbb{S}^{d-1}_+ = \mathbb{S}^{d-1} \cap \mathbb{R}^d_+$, $\alpha \mathbb{S}^{d-1}_+ = \{x \in \mathbb{R}^d_+ \mid |x| = \alpha\}$, and $[\alpha, \beta] \mathbb{S}^{d-1}_+ = \{x \in \mathbb{R}^d_+ \mid |x| \in [\alpha, \beta]\}$.

Corollary 3.1. For $0 < \alpha < \beta$ and $\sigma \in (0,1/3)$, if R is an σ -multiplicative cover of $[\alpha,\beta]\mathbb{S}^{d-1}_+$, then $|R| \geq \frac{1}{3}2^{-d}\sigma^{-d}\log(\beta/(e\alpha))$.

Proof. For any $x \in \mathbb{R}^d$, if $|x - y| < \sigma |x|$, then $|y| \in ((1 - \sigma)|x|, (1 + \sigma)|x|)$. Now define $b_k = (1 + 3\sigma)^k$ for $k \in \mathbb{N}$. Define

$$R_k = R \cap \{x \in \mathbb{R}^d_+ : |x| \in ((1 - \sigma)b_k, (1 + \sigma)b_k)\}$$

for any $k \in \mathbb{N}$. It is straightforward to check that $b_k(1+\sigma) < b_{k+1}(1-\sigma)$ since $\sigma < 1/3$. Therefore R_k s are disjoint for different ks. By construction, if $b_k \in [\alpha, \beta]$, then R_k covers $b_k \mathbb{S}^{d-1}_+$. By Lemma 3.2, we know $|R_k| \ge 2^{-d} \sigma^{-(d-1)}$ (the *multiplicative* covering of a sphere does not depend on the radius.) Thus

$$|R| \ge \sum_{k:b_k \in [\alpha,\beta]} |R_k| \ge 2^{-d} \sigma^{-(d-1)} \sum_{k:b_k \in [\alpha,\beta]} 1,$$

and we only need to count the number of b_k s that are in $[\alpha, \beta]$. Note that

$$b_k \in [\alpha, \beta] \iff k \log(1 + 3\sigma) \in [\log \alpha, \log \beta] \iff k \in [\frac{\log \alpha}{\log(1 + 3\sigma)}, \frac{\log \beta}{\log(1 + 3\sigma)}].$$

We conclude that

$$|R| \ge 2^{-d} \sigma^{-(d-1)} \left(\frac{\log \beta - \log \alpha}{\log(1 + 3\sigma)} - 1 \right)$$

relaxing $1 + 3\sigma < e^{3\sigma}$ and $1 \le 1/(3\sigma)$

$$> 2^{-d}\sigma^{-(d-1)}\left(\frac{1}{3\sigma}\log(\beta/\alpha) - \frac{1}{3\sigma}\right)$$

$$= \frac{1}{3}2^{-d}\sigma^{-d}\log(\beta/(e\alpha)),$$

which lower bounds the size of σ -multiplicative covers of $[\alpha, \beta] \mathbb{S}^{d-1}_+$.

We now finish the proof of the space lower bound of (n, d, σ) -counters.

COROLLARY 3.2. If Ω is the state space of an (n,d,σ) -counter with $\sigma < 1/3$, then

$$|\Omega| \ge \frac{1}{3} 2^{-d} (4\sigma)^{-d} \log(n/(e\sigma^{-1}d)).$$

This implies that for $n \ge e^2$, $d \ge 1$, and $\sigma < 1/16$,

$$\log_2 |\Omega| = \log_2 \log_2 n + \Omega(d \log \sigma^{-1}).$$

Proof. By Lemma 3.1, we know there exists a 4σ -multiplicative covering R of $[\sqrt{d}\sigma^{-1}, n/\sqrt{d}]$ with $|R| \leq |\Omega|$. By Corollary 3.1, we know $|R| \geq \frac{1}{3}2^{-d}(4\sigma)^{-d}\log(n/(ed\sigma^{-1}))$. This proves the first statement. For the second statement, note that

$$\log_2 |\Omega| \ge d \log_2 \sigma^{-1} - 3d - \log_2 3 + \log_2 \log(n/(e\sigma^{-1}d)).$$

If $\sigma < 1/16$ then $\log_2 \sigma^{-1} > 4$ and thus $3 \le \frac{3}{4} \log_2 \sigma^{-1}$. This implies that $d \log_2 \sigma^{-1} - 3d \ge \frac{1}{4} d \log_2 \sigma^{-1}$. Furthermore, if $\sigma^{-1} \ge n^{1/6}$ or $d \ge n^{1/6}$, then $\log_2 \log_2 n = o(d \log \sigma^{-1})$. Otherwise, we have $n/(d\sigma^{-1}) \ge n^{1/3}$. Note that $e \le n^{1/2}$ by assumption. Thus we have $\log_2 \log(n/(e\sigma^{-1}d)) \ge \log_2 \log n^{1/6} = \log_2 \log_2 n + \log_2 (\frac{1}{6} \log 2)$. We conclude that

$$\begin{split} \log_2 |\Omega| &= \log_2 \log_2 n + \Omega(d \log \sigma^{-1}) - O(1) \\ &= \log_2 \log_2 n + \Omega(d \log \sigma^{-1}), \end{split}$$

when $n \ge e^2$, $d \ge 1$, and $\sigma < 1/16$.

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A Sample Runs

For demonstration, we run the d-dimension counter defined in §2 with d=4 and memory budget $M_*=12$. The input is generated randomly at each step with probability distribution (1/2,1/4,1/8,1/8). The table below tracks the state of the algorithm whenever (U,V) is changed. Recall that x is the input count vector, the estimator is $\hat{x}=2^UV$, and the encoding of V is defined in Definition 2.1. We make the following remarks.

• When U = 0, V tracks x deterministically.

- When x is incremented, the encoding length $\psi(V)$ of V increases by at most one. When $\psi(V)$ reaches M_*+1 , the encoding length decreases by at least 1 and at most 2d during the Scale-up. when $U \geq 1$, i.e., after the first Scale-up, the encoding length is between M_*+1-2d and M_* . These are the key properties we have used in the proof of Theorem 2.1.
- Since the rates for inserting the entries are skewed, the space used to encode each entry is different. Larger entries are expected to take more space.

U	l V	estimate			1	l x				1	encoded V	
0	0000	I 0	0	0	0	 	0	0	0	0	1	
0	0100	1 0	1	0	0	i	0	1	0	0	i	10111
0	0101	1 0	1	0	1	i	0	1	0	1	İ	101101
0	1 1 1 0 1	1	1	0	1	İ	1	1	0	1	İ	0 0 0
0	2 1 0 1	1 2	1	0	1	1	2	1	0	1	Ī	1 0 0
0	2 1 0 2	1 2	1	0	2	1	2	1	0	2	Ī	1 0 1
0	2202	1 2	2	0	2	1	2	2	0	2	Ī	1 1 1
0	3 2 0 2	3	2	0	2	1	3	2	0	2	Ι	10 1 1
0	4202	4	2	0	2	1	4	2	0	2	1	11 1 1
0	4302	4	3	0	2	1	4	3	0	2	1	11 10 1
0	5302	5	3	0	2	1	5	3	0	2	1	100 10 1
0	5312	5	3	1	2	1	5	3	1	2	1	100 10 0 1
0	6 3 1 2	6	3	1	2	1	6	3	1	2	1	101 10 0 1
0	6412	6	4	1	2	1	6	4	1	2	1	101 11 0 1
0	6422	6	4	2	2	1	6	4	2	2	1	101 11 1 1
1	3 3 1 1	6	6	2	2	1	6	5	2	2	1	10 10 0 0
1	3 3 1 2	1 6	6	2	4	1	6	5	2	3	1	10 10 0 1
1	4312	8	6	2	4	1	7	5	2	3	1	11 10 0 1
1	5 3 1 2	10	6	2	4	1	9	5	2	3	1	100 10 0 1
1	6 3 1 2	12	6	2	4	1	11	5	2	3	1	101 10 0 1
1	6322	12	6	4	4	1	13	6	3	3	1	101 10 1 1
1	7 3 2 2	14	6	4	4	1	14	6	3	3	1	110 10 1 1
1	7 4 2 2	14	8	4	4	1	14	7	3	3	1	110 11 1 1
1	8 4 2 2	16	8	4	4	1	15	7	3	3	1	111 11 1 1
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2	4321	16	12	8	4	1	19	11	4	4	1	11 10 1 0
2	4421	16	16	8	4	1	19	12	4	4	1	11 11 1 0
2	4521	16	20	8	4	1	23	13	6	5	1	11 100 1 0
2	4621	16	24	8	4	1	25	14	7	5	1	11 101 1 0
3	3 3 1 0	l 24	24	8	0	1	29	14	7	5	1	10 10 0
3	4310	32	24	8	0	1	30	14	7	5		11 10 0
3	5310	40	24	8	0	1	31	14	7	5		100 10 0
3	5410	40	32	8	0	1	36	16	7	6		100 11 0
3	5420	40	32	16	0	1	45	20	9	7		100 11 1
3	6420	48	32	16	0	1	50	25	11	8		101 11 1
	7 4 2 0	56	32	16	0	1	52	28	12	8		110 11 1
	7 4 3 0	56	32	24	0	I	52	28	13	8		110 11 10
	8 4 3 0	l 64	32	24	0	I	57	33	14	12		111 11 10
	4210	64	32	16	0	I	64	34	15	14		11 1 0
	5 2 1 0	l 80	32	16	0		76	43	17	17	I	100 1 0
	6210	96	32	16	0		84	46	19	18	I	101 1 0
	7 2 1 0	112	32	16	0		90	49	20	18	I	110 1 0
	7 3 1 0	112	48	16	0	I	90	50	20	18		110 10 0
4	8310	128	48	16	0	1	93	51	22	19	1	111 10 0

4	9310	144 4	8 16	0	1	94	53	22	19	ı	1000 10 0
4			8 16		i	125	63	25	25	i	1001 10 0
4			4 16		i	134	69	28	29	i	1001 11 0
5			4 32		Ĺ	137	69	28	30	Ĺ	100 1 0 0
5			6 32		ĺ	139	73	29	30	ĺ	100 10 0 0
5	5411	160 12	8 32	32	- 1	154	76	32	30	-	100 11 0 0
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5	7 4 2 1	224 12	8 64	32	- 1	236	105	55	46	- 1	110 11 1 0
5	8 4 2 1	256 12	8 64	32	- 1	277	118	62	55	- 1	111 11 1 0
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6	4211	l 256 12	8 64	64		315	133	70	64		11 1 0 0
6	5211	320 12	8 64	64		322	139	73	65		100 1 0 0
6	6211	384 12	8 64	64		345	154	79	66		101 1 0 0
6	6212	384 12	8 64	128	- 1	374	176	88	76	- 1	101 1 0 1
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6	9212	576 12	8 64	128		478	225	120	107		1000 1 0 1
7		512 12				593	300	134	136		11 0 0
7		640 12				599	302	139	138		100 0 0
7		640 12				634	321	146	146		100 0 0 0
7		768 12				686	351	153	156		101 0 0 0
7		768 25				687	353	155	157		101 1 0 0
7		768 38			-	732	377	170	174		101 10 0 0
7		896 38				785	419	183	189		110 10 0 0
7		896 51				869	464	203	216		110 11 0 0
7	8411	1024 51				969	522	235	236		111 11 0 0
7	8412	1024 51				994	536	244	240		111 11 0 1
8 8	4 2 0 1 4 3 0 1	1024 51 1024 76			 	1058 1102	570 596	259 269	252 262		11 1 0 11 10 0
8	5301	1024 76			i	1138	623	282	274	i	100 10 0
8	6301	1536 76			i	1217	660	304	295	i	101 10 0
8	7 3 0 1	1792 76			i	1222	665	307	296	i	110 10 0
8	7 4 0 1	1792 102			i	1344	718	339	321	i	110 11 0
8	7501	1792 128			i	1521	812	377	366	i	110 100 0
8	7 5 0 2	1792 128			i	1581	850	393	384	i	110 100 1
9	3 3 0 1	1536 153			i	1713	907	427	423	i	10 10 0
9	4301	2048 153			İ	1743	922	430	427	İ	11 10 0
9	5301	2560 153			Ì	1755	930	434	434	İ	100 10 0
9	5401	2560 204	8 0	512		1849	987	460	453	-	100 11 0
9	6401	3072 204	8 0	512		2215	1195	572	548	-	101 11 0
9	6411	3072 204	8 512	512		2451	1293	630	607		101 11 0 0
9	7 4 1 1	3584 204	8 512	512		2481	1311	640	615		110 11 0 0
9	8 4 1 1	4096 204	8 512	512		2663	1396	691	660		111 11 0 0
9	8 4 2 1	4096 204	8 1024	512		2751	1433	719	686		111 11 1 0
10	4211	l 4096 204					1458	733	695		11 1 0 0
10	4221	4096 204					1781	876	836		11 1 1 0
10	4222	4096 204					2071		979		11 1 1 1
10	4322	4096 307	2 2048	2048	I	4217	2145	1071	1021	ı	11 10 1 1