# Message-Passing Algorithms and Improved LP Decoding<sup>\*</sup>

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## ABSTRACT

Linear programming decoding for low-density parity check codes (and related domains such as compressed sensing) has received increased attention over recent years because of its practical performance —coming close to that of iterative decoding algorithms— and its amenability to finite-blocklength analysis. Several works starting with the work of Feldman et al. showed how to analyze LP decoding using properties of expander graphs. This line of analysis works for only low error rates, about a couple of orders of magnitude lower than the empirically observed performance. It is possible to do better for the case of random noise, as shown by Daskalakis et al. and Koetter and Vontobel.

Building on work of Koetter and Vontobel, we obtain a novel understanding of LP decoding, which allows us to establish a 0.05-fraction of correctable errors for rate-1/2codes; this comes very close to the performance of iterative decoders and is significantly higher than the best previously noted correctable bit error rate for LP decoding. Unlike other techniques, our analysis directly works with the primal linear program and exploits an explicit connection between LP decoding and message passing algorithms.

An interesting byproduct of our method is a notion of a "locally optimal" solution that we show to always be globally optimal (i.e., it is the nearest codeword). Such a solution can in fact be found in near-linear time by a "re-weighted" version of the min-sum algorithm, obviating the need for linear programming. Our analysis implies, in particular, that this re-weighted version of the min-sum decoder corrects up to a 0.05-fraction of errors.

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Algorithms, Theory

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LDPC codes, LP decoding, min-sum algorithm, messagepassing algorithms

# 1. INTRODUCTION

Low density parity-check (LDPC) codes are linear codes over GF(2) whose constraint graph is sparse. They were introduced and analyzed by Gallager [10] in a paper that was forgotten for several decades and recalled again only in the 1990s. Sipser and Spielman [17] studied a subclass of these codes in which the constraint graph has good expansion properties. For these expander codes, they showed that a very simple *bit-flipping* strategy, originally suggested by Gallager, corrects efficiently an  $\Omega(1)$  fraction of (worst-case) errors, though the actual constants were quite weak. (Myriad extensions of expander codes have been studied but will not be relevant here.)

Meanwhile, researchers in information theory rediscovered Gallager's idea and began to view the decoding problem for LDPC codes as an example of *Maximum a posteriori* (MAP) estimation in factor graphs (a notion that also became popular in machine learning). Various iterative message-passing algorithms—two popular ones being belief propagation (BP) and the *min-sum algorithm*—were found to empirically yield excellent decoding performance. A survey of LDPC codes and decoding algorithms appears in [16, 14].

In a seminal paper [14], Richardson and Urbanke, aided with some computer calculations, were able to establish that a BP-like algorithm can decode with high probability a (3, 6)regular LDPC code on a binary symmetric channel with noise rate up to 0.084. This is the best bound known for any decoding algorithm, and is not far from the empirically observed performance of BP and the information theoretic limit of roughly 0.11.

Our paper concerns the linear programming (LP) decoding method. This method was introduced by Feldman, Wainwright and Karger [8], in a paper that only establishes a sub-linear number of correctable errors, but also notes that the empirical performance of LP decoding is similar to that of message-passing algorithms. A subsequent paper of Feldman et al. [7] showed that the method corrects  $\Omega(1)$ -fraction of (worst-case) errors for expander codes. The proof consists of constructing a dual solution, inspired by Sipser and Spielman's analysis, and yields similar bounds: a tolerance of

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adversarial noise rate up to 0.00017. Note that the advantage of LP decoding over message-passing decoders is that in the case of decoding success, the linear program provides a certificate that the output is indeed the nearest codeword.

Unfortunately, it has remained difficult to improve the analysis of LP decoding to establish bounds closer to the empirically observed performance. Daskalakis et al. [5] were able to show a tolerance of noise up to 0.002 —an order of magnitude better than the bounds of Feldman et al. [7] but still 40 times lower than the Richardson–Urbanke [14] bound of 0.084 for belief propagation. Their proof constructs a more intricate dual LP solution than Feldman et al.'s, but it is still based on expansion arguments. (Note: All the bounds in the paper are quoted for (regular) rate 1/2 codes.)

Intuitively, the main reason for the small bit error rates in the above analyses of LP decoding was that these analyses were close in spirit to the Sipser and Spielman expansionbased approach. By contrast the Richardson–Urbanke style analysis of message passing algorithms relies upon the *high* girth of the graph defining the code (specifically, the fact that high-girth graphs look locally like trees).

Nevertheless, it remained unclear how to bring girth-based arguments into the context of LP decoding. In a recent paper, Koetter and Vontobel [13] achieved this. Their key idea was to use the min-sum algorithm rather than Belief-Propagation (which uses highly nonlinear operations). They showed how to transform the messages exchanged in the min-sum algorithm into an intricate dual solution. (Their construction was inspired by the Gauss–Seidel method to solve convex programs.) Though they did not report any numbers in their paper, our calculations show that their analysis of LP decoding allows (3, 6)-regular codes to tolerate random noise rate 0.01 —a factor of 5 improvement over Daskalakis et al. [5].

In this paper we present an improvement of the noise rate by another factor of 5 to 0.05, coming very close to the performance of BP. The key ingredient in our proof is a new approach to analyzing LP decoding. Instead of trying to construct a dual solution as in all the previous papers, we give a direct analysis of the primal linear program. (This also answers an open question of Feldman et al. regarding whether a primal-only analysis is possible.) At its heart, the proof relies on the fact that the LP relaxation is tight for trees. We use this to show that an LP solution can be decomposed into a distribution over codewords for every tree-like neighborhood of G so that these distributions are consistent in overlapping neighborhoods; the type of consistency that we use is inspired by hierarchies of LP relaxations, such as the Sherali–Adams hierarchy [15]. We use our decomposition to define a criterion for certifying the optimality of a codeword in the right circumstances (Theorem 2), which is quite interesting on its own right. If the certificate exists, it can be found by a simple message-passing algorithm, and if it exists, then LP decoding works (Theorem 4). The first such certificate was described in [13]; ours is more general and therefore occurs with high probability for much larger noise rates (Theorems 1, 5). We note that prior to [13] no other analyses led to message-passing algorithms that *certify* the correctness of their answer.

As for the probability with which such a certificate exists, our calculation consists of reducing the whole problem to the study of a min-sum process on a finite tree (Definition 4), which is even amenable to analytic calculation by hand, as done for error rate up to 0.0247 (see Section 6.1). This consists of tracing the Laplace transform of the messages exchanged by the min-sum process, as these messages move upwards on the tree. We believe that this idea of recursing the Laplace transform, rather than the density functions, of the messages is interesting on its own right and could be useful in other settings. In our setting it is rather effective in handling the min operators, which we cannot handle analytically if we trace the density functions of the messages.

Combining our analytic bounds with a MATLAB calculation, we can accommodate noise rate up to 0.05 (see Section 6.1). The method seems to break down beyond 0.05, suggesting that getting to 0.084 would require new ideas. We note that our analysis does not require expansion, only high enough girth (a lower bound of  $\Omega(\log \log n)$  on the girth is sufficient to make the probability of decoding error inverse polynomial in the blocklength n). Perhaps the right idea to go beyond 0.05 is to marry high-girth and expansion-based arguments, an avenue worth exploring.

An interesting byproduct of our technique is establishing that a certain re-weighted version of the min-sum decoder corrects up to a 0.05-fraction of errors with high probability over the binary symmetric channel for code-rate 1/2. To the best of our knowledge, the bound of 0.05 is the best known for re-weighted min-sum decoders over BSC for codes of rate 1/2(c.f. [18, 9, 3, 4, 2]). As compared to the 0.084 bound for BP, ours has the advantage that with high probability the nearest codeword can be certified to be correct.

We note that our method, being primal-only, is relatively clean —and in our opinion, easier to understand (apart maybe from the probabilistic calculation) than previous analyses. We also suspect that the idea of expressing primal solutions in terms of local tree assignments may be of wider use in applications that use LP decoding techniques and random graphs. We are currently exploring connections to compressed sensing. Candès and Tao [1] (independently of Feldman et al. though somewhat later), as part of work on *compressed* sensing, arrived at linear programming as a promising tool for a variety of reconstruction problems in compressed sensing, which include decoding random linear codes over the reals (these are not LDPC since the constraint graph is non-sparse). Recent work such as [11, 12] makes explicit the connection between Sipser-Spielman type decoding of LDPC codes and compressed sensing using sparse matrices.

#### Our Main Result.

THEOREM 1. Let  $p \leq 0.05$  and let  $x \in \{0,1\}^n$  be a codeword of the low-density parity check code defined by a (3,6)regular bipartite graph with  $\Omega(\log n)$  girth. Suppose that  $y \in \{0,1\}^n$  is obtained from x by flipping every bit independently with probability p. Then, with probability at least  $1 - \exp(-n^{\gamma})$  for some constant  $\gamma > 0$ ,

- the codeword x is the unique optimal solution to the LP decoder of Feldman et al. [8] (see LP (2) in Section 2),
- 2. a simple message-passing (dynamic programming) algorithm running in time  $O(n \log n)$  can find x and certify that it is the nearest codeword to y.

For LDPC codes defined by general  $(d_L, d_R)$ -regular graphs, we have the same conclusion whenever  $d_L$ ,  $d_R$ , and p satisfy the condition

$$\sqrt{p} \left(1 - (1-p)^{d_R-1}\right)^{\frac{d_L-2}{2}} (1-p)^{\frac{(d_R-1)(d_L-2)}{2} + \frac{1}{2}} < \frac{1}{(d_R-1) \cdot 2^{d_L-1}} . \quad (1)$$

Remark 1. If we are content with a decoding success probability of 1 - 1/poly(n), then  $\Omega(\log \log n)$  girth is sufficient for the results in the previous theorem. The running time is reduced to  $O(n \log \log n)$ .

## 2. PRELIMINARIES

#### *Low-density parity check codes.*

Let G be a simple bipartite graph with bipartition  $(V_L, V_R)$ , left degree  $d_L$ , and right degree  $d_R$ . Let n be the number of left vertices, and m be the number of right vertices. We will assume that  $V_L$  is the set  $[n] \stackrel{\text{def}}{=} \{1, \ldots, n\}$ . For two vertices u and v of G, we let d(u, v) be the distance of u and v in G. We denote by N(u) the set of neighbors of u. Similarly,  $N^t(u)$  is the set of vertices at distance t from u. We denote by  $N^{\leq t}(u)$  or B(u, t) the set of vertices at distance at most t from u.

The *parity-check code* defined by G is the set of all 0/1 assignments to the left vertices such that every right vertex has an even number of neighbors with value 1,

$$\mathcal{C}(G) \stackrel{\text{def}}{=} \{x \in \{0,1\}^n \mid \sum_{i \in N(j)} x_i \equiv 0 \mod 2, \text{ for all } j \in V_R\}.$$

The elements of  $\mathcal{C}(G)$  are called *codewords*. Note that if we allow general graphs then any linear code can be realized as parity check code  $\mathcal{C}(G)$  for some graph G. In this paper, we will only deal with sparse (low-density) graphs, that is, the degrees  $d_L$  and  $d_R$  will be constants.

In the following, we refer to the vertices in  $V_L$  and  $V_R$  as variable nodes and check nodes, respectively.

#### LP Decoding.

In the nearest codeword problem for the code  $\mathcal{C}(G)$ , we are given a vector  $y \in \{0,1\}^n$  and the goal is to find a codeword  $x \in \mathcal{C}(G)$  so as to minimize the Hamming distance  $||x - y||_1$ .

In [6, 8], Feldman et al. introduced the following LP relaxation for this problem:

$$Minimize ||x - y||_1 (2)$$

subject to 
$$x \in \bigcap_{j \in V_R} \operatorname{Conv} \mathcal{C}_j$$
, (3)

where Conv X denotes the convex hull of a set X of bit vectors, and  $C_j$  is the set of bit vectors satisfying constraint  $j \in V_R$ ,

$$\mathcal{C}_j \stackrel{\text{def}}{=} \left\{ x \in \{0,1\}^n \mid \sum_{i \in N(j)} x_i \equiv 0 \mod 2 \right\}$$

We call  $x \in [0, 1]^n$  an *LP* solution if it satisfies (3). We say x is an optimal *LP* solution given y if x is an LP solution that achieves the minimum distance (2) to y. An optimal LP solution can be computed in time polynomial in n. In this paper, we are interested to find conditions under which the solution of this LP coincides with the nearest codeword to y.

Before concluding this section, we note that  $||x - y||_1$  is an affine linear function of  $x \in [0, 1]^n$  for any fixed  $y \in \{0, 1\}^n$ , since  $||x - y||_1 = ||y||_1 + \sum_{i=1}^n (-1)^{y_i} x_i$ , for all  $x \in [0, 1]^n$ .

## 3. CERTIFYING THE NEAREST CODE-WORD

This section considers the following question: Given a codeword  $x \in C(G)$  and a vector  $y \in \{0,1\}^n$ , how can we certify efficiently that x is the nearest codeword to y? We present a certificate based on local checks that is inspired by and generalizes the key idea in the calculation of Koetter and Vontobel [13]. The motivation for this generalization is that it allows certification/decoding in the presence of much higher noise. Our proof that this certificate works is also quite different. It is designed to easily carry over to prove that x is also the unique fractional solution to the LP.

Throughout this section and the next, y is the received word and x is a codeword which we are trying to certify as the nearest codeword to y. We will be considering assignments to neighborhoods  $N^{\leq 2T}(i_0)$  where  $i_0 \in V_L$  and  $T < \frac{1}{4} \operatorname{girth}(G)$ . Thus the induced graph on  $N^{\leq 2T}(i_0)$  is a tree with degrees  $d_L, d_R$  respectively at even and odd levels (the level of a node is its distance to  $i_0$ ). Note that the variable nodes in  $N^{\leq 2T}(i_0)$  have even distance to  $i_0$  and the check nodes have odd distance.

The motivation for considering local neighborhoods comes from known analyses of message-passing algorithms. Such algorithms are local in the following sense: after  $t < \frac{1}{4} \operatorname{girth}(G)$ iterations, the value computed for the variable *i* is a "guess" for  $x_i$  given the information in the neighborhood N(i, 2t); in this sense, after *t* rounds message passing algorithms compute a "locally optimal" solution. Several notions of "local optimality" were implicit in the algorithms of [18, 9, 3, 4, 2]. Our notion of local optimality generalizes the notions used in all of these papers, and our interest centers in showing that local optimality implies global optimality.

Our notion of local optimality is given in Definition 2 and requires the following definition generalizing Wiberg [18].

Definition 1. (Minimal Local Deviation) An assignment  $\beta \in \{0, 1\}^n$  is a valid deviation of depth T at  $i_0 \in V_L$  or, in short, a T-local deviation at  $i_0$ , if  $\beta_{i_0} = 1$  and  $\beta$  satisfies all parity checks in  $B(i_0, 2T)$ ,

$$\forall j \in V_R \cap B(i_0, 2T): \quad \sum_{i \in N(j)} \beta_i \equiv 0 \bmod 2$$

(Notice that  $\beta$  need not be a codeword since we do not insist that the check nodes beyond level 2T from  $i_0$  are satisfied.)

A *T*-local deviation  $\beta$  at  $i_0$  is minimal if  $\beta_i = 0$  for every  $i \notin B(i_0, 2T)$ , and every check node j in  $B(i_0, 2T)$  has at most two neighbors with value 1 in  $\beta$ . Note that a minimal *T*-local deviation at  $i_0$  can be seen as a subtree of  $B(i_0, 2T)$  of height 2*T* rooted at  $i_0$ , where every variable node has full degree and every check node has degree 2. We will refer to such trees as skinny trees.

An assignment  $\beta \in \{0, 1\}^n$  is a minimal *T*-local deviation if it is a minimal *T*-local deviation at some  $i_0$ . Note that given  $\beta$  there is a unique such  $i_0 \stackrel{\text{def}}{=} \operatorname{root}(\beta)$ .

If  $w = (w_1, \ldots, w_T) \in [0, 1]^T$  is a weight vector and  $\beta$  is a minimal *T*-local deviation, then  $\beta^{(w)}$  denotes the *w*-weighted deviation

$$\beta_i^{(w)} = \begin{cases} w_t \beta_i & \text{if } d(\operatorname{root}(\beta), i) = 2t \text{ and } 1 \leqslant t \leqslant T, \\ 0 & \text{otherwise.} \end{cases}$$

(End of Definition 1.)

For two vectors  $u, v \in \{0, 1\}^n$ , we denote by  $u \oplus v$  the coordinate-wise sum of u and v modulo 2. We extend this notation to fractional vectors in the following way: If  $u \in \{0, 1\}^n$  and  $\bar{v} \in [0, 1]^n$ , then  $u \oplus \bar{v} \in [0, 1]^n$  denotes the vector with  $i^{\text{th}}$  coordinate  $|u_i - \bar{v}_i|$ . Note that, for a fixed vector  $u \in \{0, 1\}^n$ ,  $u \oplus \bar{v}$  is affine linear in  $\bar{v}$ . Hence, for any distribution over vectors  $v \in \{0, 1\}^n$  and a fixed bit vector u, we have  $\mathbb{E} u \oplus v = u \oplus (\mathbb{E} v)$ .

Definition 2. (Local optimality) A codeword  $x \in \{0,1\}^n$ is (T, w)-locally optimal for  $y \in \{0,1\}^n$  if for all minimal *T*-local deviations  $\beta$ ,

$$||x \oplus \beta^{(w)} - y||_1 > ||x - y||_1$$

Since the local neighborhood at node  $i_0$  is tree-like, the minimal T-local deviation  $\beta$  at  $i_0$  that minimizes  $||x \oplus \beta^{(w)} - y||_1$  can be computed by a simple dynamic programing algorithm in time proportional to the size of the neighborhood. In fact, we can do the computations for all the  $i_0$ 's simultaneously with dynamic programming to achieve a total running time of  $O(T \cdot n)$  (which is nearly linear since we are going to choose  $T = O(\log n)$  or  $T = O(\log \log n)$  later in the proof).

Note that if x is locally optimal then this is some intuitive evidence that x is the nearest codeword, since changing x in just one variable  $i_0$  seems to cause the distance from y to increase in the immediate neighborhood of  $i_0$ . Koetter and Vontobel [13] make this intuition precise for w = 1, in which case they show that a locally optimal x is also globally optimal, that is, the nearest codeword to y. (In fact, all previous theoretical analyses of message passing algorithms have some notion of local optimality but none of them were known to imply global optimality.) Our proof works for general w.

THEOREM 2. Let  $T < \frac{1}{4} \operatorname{girth}(G)$  and  $w = (w_1, \ldots, w_T)$ be a non-negative weight vector. If x is a (T, w)-locally optimal codeword for  $y \in \{0, 1\}^n$ , then x is also the unique nearest codeword for y.

Theorem 2 implies that we can certify that x is the nearest codeword for y by verifying the local optimality condition. It raises two questions:

- 1. How can we find a locally optimal codeword if it exists?
- 2. What is the chance that the nearest codeword satisfies the local optimality condition?

The first question has been studied in the context of messagepassing algorithms. For w = 1, Wiberg [18] showed that the well-known min-sum decoding algorithm can find locally optimal codewords. Other specific weight functions were suggested and analyzed in several works [9, 3, 4, 2]. We show how to to make min-sum decoding work for arbitrary weight vectors (which allows much more freedom in deriving analytic error bounds). We refer to the full version of this paper for further details.

THEOREM 3. Let  $T < \frac{1}{4} \operatorname{girth}(G)$  and  $w = (w_1, \ldots, w_T)$ be a non-negative weight vector. Suppose that x is a (T, w)locally optimal codeword for  $y \in \{0, 1\}^n$ . Then the wreweighted min-sum algorithm on input y computes x in T iterations. Since the focus of this paper is on LP decoding, we next address whether LP decoding can find locally optimal solutions. To this end we extend Theorem 2 in order to show that a locally optimal solution is not only the nearest codeword to y but also the unique optimal LP solution given y. For the case w = 1, this was also established by Koetter and Vontobel [13].

THEOREM 4. Let  $T < \frac{1}{4} \operatorname{girth}(G)$  and  $w = (w_1, \ldots, w_T)$ be a non-negative weight vector. Suppose that x is a (T, w)locally optimal codeword for  $y \in \{0, 1\}^n$ . Then x is also the unique optimal LP solution given y.

The proof for w = 1 in [13] proceeded by constructing an appropriate dual solution in an iterative manner. Our proof in Section 5 yields a more general result, and is completely different, since it only looks at the primal LP.

Now we address the second question regarding the probability that the nearest codeword is locally optimal. (Showing higher probabilities for this event was the main motivation for introducing general weight vectors w.) We prove the following theorem in Section 6.

THEOREM 5. Let G be a  $(d_L, d_R)$ -regular bipartite graph and  $T < \frac{1}{4} \operatorname{girth}(G)$ . Let also  $p \in (0, 1)$  and  $x \in \{0, 1\}^n$  be a codeword in  $\mathcal{C}(G)$ . Suppose that y is obtained from x by flipping every bit independently with probability p.

#### 1. If $d_L$ , $d_R$ , and p satisfy the condition

$$\min_{t>0} \left\{ \left( (1-p) \ e^{-t} + p \ e^{t} \right) \\ \cdot \left( (1-p)^{d_{R}-1} \ e^{-t} + (1-(1-p)^{d_{R}-1}) \ e^{t} \right)^{d_{L}-2} \right\} \\ < \frac{1}{d_{R}-1}, \quad (4)$$

then x is (T, 1)-locally optimal with probability at least  $1 - n \cdot c^{-(d_L - 1)^T}$  for some c > 1. For  $(d_L, d_R) = (3, 6)$ , Condition (4) is satisfied whenever  $p \leq 0.02$ .

2. If  $d_L$ ,  $d_R$ , and p satisfy the condition

$$\sqrt{p} \left(1 - (1-p)^{d_R-1}\right)^{\frac{d_L-2}{2}} (1-p)^{\frac{(d_R-1)(d_L-2)}{2} + \frac{1}{2}} 
< \frac{1}{(d_R-1) \cdot 2^{d_L-1}}, \quad (5)$$

then there exists a weight vector  $w \in [0, 1]^T$  such that x is (T, w)-locally optimal with probability at least  $1 - n \cdot c^{-(d_L - 1)^T}$  for some constant c > 1. For  $(d_L, d_R) = (3, 6)$ , Condition (5) is satisfied whenever  $p \leq 0.0247$ .

3. If  $(d_L, d_R) = (3, 6)$  and  $p \leq 0.05$ , then x is (T, w)locally optimal with probability at least  $1 - n \cdot c^{-2^T}$  for some weight vector w and some constant c > 1.

Given Theorems 2, 3, 4, and 5 we can obtain the proof of Theorem 1 as follows: take  $T = \Theta(\log n) < \frac{1}{4} \operatorname{girth}(G)$ ; from Theorem 5, there exists a weight vector w such that, with probability at least  $1 - \exp(-n^{\gamma})$  (for some constant  $\gamma$  depending on the leading constant in front of  $\log n$  in the choice of T), the codeword x is (T, w)-locally optimal. From Theorem 4 it follows then that x is the unique optimal LP solution given y. Also, from Theorem 3, it follows that x can be found by the w-reweighted min-sum algorithm in  $T = O(\log n)$  iterations, so time  $O(n \log n)$  overall, and by Theorem 2 it can be certified that x is the nearest codeword to y. (If the girth is  $\Omega(\log \log n)$  then decoding still succeeds with probability  $1 - 1/\operatorname{poly}(n)$ .)

# 4. LOCAL OPTIMALITY IMPLIES GLOBAL OPTIMALITY

#### *Proof of Theorem* **2***.*

In this section  $y \in \{0, 1\}^n$  is the received word,  $x \in \{0, 1\}^n$  is a locally optimal codeword in  $\mathcal{C}(G)$ , and  $x' \in \{0, 1\}^n$  is a codeword different from x. We wish to show  $||x' - y||_1 > ||x - y||_1$ .

The following lemma is the key to our proof of Theorem 2.

LEMMA 1. Let  $T < \frac{1}{4}$  girth(G). Then, for every codeword  $z \neq 0$ , there exists a distribution over minimal T-local deviations  $\beta$  such that for every weight vector  $w \in [0, 1]^T$ ,

$$\mathbb{E}\,\beta^{(w)} = \alpha z\,,$$

where  $\alpha \in [0, 1]$  is some scaling factor.

Before proving the lemma, let us first see how we can finish the proof of Theorem 2 using such a distribution over minimal local deviations.

PROOF OF THEOREM 2. Let x be a (T, w)-locally optimal codeword for  $y \in \{0, 1\}^n$ . We want to show that for every codeword  $x' \neq x$ , the distance to y increases, that is,  $||x - y||_1 < ||x' - y||_1$ . The main idea is to observe that  $z = x \oplus x'$  is also a codeword, and hence by Lemma 1, there exists a distribution over minimal T-local deviations  $\beta$  such that  $\mathbb{E} \beta^{(w)} = \alpha z$  for the codeword  $z = x \oplus x'$ . Now it is easy to complete the proof using local optimality of x. Let  $f: [0,1]^n \to \mathbb{R}$  be the affine linear function  $f(u) = ||x \oplus u - y||_1 = ||x - y||_1 + \sum_{i=1}^n (-1)^{x_i + y_i} u_i$ . Now,

$$\begin{split} \|x - y\|_1 &< \mathbb{E} \|x \oplus \beta^{(w)} - y\|_1 \qquad \text{(by local optimality of } x) \\ &= \|x \oplus (\alpha z) - y\|_1 \qquad \text{(affine linearity of } f) \\ &= \alpha \|x' - y\|_1 + (1 - \alpha) \|x - y\|_1 \quad \text{(aff. lin. of } f) \,, \end{split}$$

which implies  $||x - y||_1 < ||x' - y||_1$  as desired.  $\Box$ 

## 4.1 Proof of Lemma 1

Constructing Distributions over Minimal Local Deviations for Codewords.

Let  $z \in \{0,1\}^n$  be a codeword. We want to construct a distribution over minimal local deviations such that the mean of the distribution is proportional to z.

For every variable node  $i \in V_L$  with  $z_i \neq 0$ , we define a distribution over subtrees  $\tau_i$  of G of height 2T rooted at i: The idea is that we grow  $\tau_i$  minimally and randomly inside the non-zeros of z starting from the variable node i. Consider the neighborhood  $N^{\leq 2T}(i)$  and direct the edges away from the root i. Remove all variable nodes in this neighborhood with value 0 in z. Remove now the vertices that are no longer reachable from i. In the remaining tree (rooted at i), pick a random subtree  $\tau_i$  with full out-degree at variable nodes and out-degree 1 at check nodes.

Suppose now that we choose such a tree  $\tau_i$  for all *i* with  $z_i \neq 0$ , such that these trees are mutually independent. Independently from the choices of the trees, we also choose

 $i_0$  uniformly at random from the support of z (that is, we pick  $i_0$  with probability  $z_{i_0}/||z||_1$ ), and define  $\beta$  as

$$\beta_i = \begin{cases} 1 & \text{if } i \in \tau_{i_0} ,\\ 0 & \text{otherwise.} \end{cases}$$

We denote by  $\mathbb{P}$  the joint probability measure over the trees  $\{\tau_i\}_{i:z_i\neq 0}$ , the variable node  $i_0$  and the assignment  $\beta$ . Before concluding the proof of Lemma 1, we make a few observations about the random subtrees  $\tau_i$ . First, the number of nodes at level  $2t \leq 2T$  of any tree  $\tau_i$  is always exactly  $d_L(d_L - 1)^{t-1}$  (the root has out-degree  $d_L$ ). Second, for any two variable nodes i, i' with  $z_i = z_{i'} = 1$  the above process treats them symmetrically:

$$\mathbb{P}\left\{i' \in \tau_i\right\} = \mathbb{P}\left\{i \in \tau_{i'}\right\}.$$
(6)

If d(i, i') > 2T, then both of the probabilities are 0 (since the height of the trees is 2T). Otherwise, the nodes i, i' are connected by a unique path of length  $\leq 2T$ , say  $(i, j_0, i_1, j_1, \dots, i_{t-1}, j_{t-1}, i')$ . If there exists some variable node  $i_{\ell}, \ell \in \{1, \dots, t-1\}$ , in this path with  $z_{i_{\ell}} = 0$  then both of the probabilities are zero. Otherwise, let  $d_r = \sum_{i \in N(j_r)} z_j$ be the number of neighbors of  $j_r$  that are in the support of z. Then both of the probabilities in (6) are equal to

$$rac{1}{(d_0-1)\cdots(d_{t-1}-1)}$$
 .

Armed with these observations, we can analyze our distribution over minimal local deviations  $\beta$ : If  $z_i = 0$ , then  $\beta_i = 0$ with probability 1. Hence, we may assume  $z_i = 1$ . Then,

$$\mathbb{E}\beta_{i}^{(w)} = \sum_{t=1}^{T} w_{t} \sum_{\substack{i' \in N^{2t}(i) \\ z_{i'}=1}} \mathbb{P}\left\{i_{0}=i'\right\} \mathbb{P}\left\{i \in \tau_{i'}\right\}$$

$$\stackrel{(6)}{=} \sum_{t=1}^{T} w_{t} \sum_{\substack{i' \in N^{2t}(i) \\ z_{i'}=1}} \frac{1}{\|z\|_{1}} \mathbb{P}\left\{i' \in \tau_{i}\right\}$$

$$= \frac{1}{\|z\|_{1}} \sum_{t=1}^{T} w_{t} \mathbb{E}\left[\tau_{i} \cap N^{2t}(i)\right]$$

$$= \frac{1}{\|z\|_{1}} \sum_{t=1}^{T} w_{t} \cdot d_{L}(d_{L}-1)^{t-1}$$

Therefore, we have the desired conclusion  $\mathbb{E}\beta^{(w)} = \alpha z$  with  $\alpha = \sum_{t=1}^{T} w_t \cdot d_L (d_L - 1)^{t-1} / ||z||_1$ .  $\Box$ 

# 5. LOCAL OPTIMALITY IMPLIES LP OP-TIMALITY

#### Proof of Theorem 4.

Let  $x \in \{0,1\}^n$  be a codeword in  $\mathcal{C}(G)$  and let  $y \in \{0,1\}^n$ . The following lemma is completely analogous to Lemma 1, and follows from the fact that LP solutions look locally like distributions over codewords.

LEMMA 2. Let  $T < \frac{1}{4} \operatorname{girth}(G)$  and  $w \in [0, 1]^T$ . Then for every non-zero LP solution  $z \in [0, 1]^n$ , there exists a distribution over minimal T-local deviations  $\beta$  such that

$$\mathbb{E}\beta^{(w)} = \alpha z \,,$$

where  $\alpha_t \stackrel{\text{def}}{=} \frac{d_L (d_L - 1)^{t-1}}{\|x - x'\|_1}$ .

Using a distribution over minimal local deviations from Lemma 2, we can prove Theorem 4 in almost the same way as Theorem 2 in the previous section. The only additional ingredient is the following simple property of LP solutions. Recall that for  $x \in \{0,1\}^n$  and  $x' \in [0,1]^n$ , we denote by  $x \oplus x'$  the vector whose *i*th coordinate is  $|x_i - x'_i|$ . The next lemma is straightforward using the observation that the defining property of an LP solution (specifically, (3)) is that locally (for every check node) it can be viewed as convex combinations of even-parity vectors.

LEMMA 3. Let x be a codeword and x' be LP solutions (i.e., satisfy (3)). Then  $x \oplus x'$  is also an LP solution.

Now we can prove the main theorem.

PROOF OF THEOREM 4. Let x be a (T, w)-locally optimal codeword for  $y \in \{0, 1\}^n$ . We want to show that for every LP solution  $x' \neq x$ , the distance to y increases, that is,  $||x-y||_1 <$  $||x'-y||_1$ . By Lemma 2, there exists a distribution over minimal T-local deviations  $\beta$  such that  $\mathbb{E} \beta^{(w)} = \alpha z$  for the LP solution  $z = x \oplus x'$ . Let  $f : [0, 1]^n \to \mathbb{R}$  be the affine linear function  $f(u) = ||x \oplus u - y||_1 = ||x - y||_1 + \sum_{i=1}^n (-1)^{x_i+y_i} u_i$ . Now,

$$\begin{split} \|x - y\|_1 &< \mathbb{E} \|x \oplus \beta^{(w)} - y\|_1 \qquad \text{(by local optimality of } x) \\ &= \|x \oplus (\alpha z) - y\|_1 \qquad \text{(affine linearity of } f) \\ &= \alpha \|x' - y\|_1 + (1 - \alpha) \|x - y\|_1 \quad \text{(aff. lin. of } f) \,, \end{split}$$

which implies  $||x - y||_1 < ||x' - y||_1$  as desired.  $\Box$ 

## 5.1 Proof of Lemma 2

Constructing Distributions over Minimal Local Deviations for LP Solutions.

Let  $z \in [0,1]^n$  be a non-zero LP solution. The proof of the current lemma is very similar to the proof of Lemma 1 (the integral case). The following lemma is essentially the only additional ingredient of the proof.

LEMMA 4. For every  $j \in V_R$ , we can find a function  $\rho_j: V_L \times V_L \to \mathbb{R}_+$  such that

1. for every neighbor  $i \in N(j)$ 

$$z_i = \sum_{i' \in N(j) \setminus \{i\}} \rho_j(i, i')$$

2. for any two neighbors  $i, i' \in N(j)$ ,

$$\rho_j(i,i') = \rho_j(i',i) \,.$$

PROOF. Since z is an LP solution, it is a convex combination of assignments in  $C_j = \{\gamma \in \{0,1\}^n \mid \sum_{i \in N(j)} \gamma_i \equiv 0 \mod 2\}$ . Hence, there are multipliers  $\alpha_{\gamma} \ge 0$  with  $\sum_{\gamma \in C_j} \alpha_{\gamma} = 1$  such that  $z = \sum_{\gamma \in \{0,1\}^n} \alpha_{\gamma} \gamma$ . Now we can define  $\rho_j(i,i')$  as

$$\rho_j(i,i') = \sum_{\substack{\gamma \in \mathcal{C}_j \\ \gamma_i = 1, \gamma_{i'} = 1}} \alpha_\gamma \frac{\gamma_i \gamma_{i'}}{\sum_{i'' \in N(j) \setminus \{i\}} \gamma_{i''}}$$

The second property (symmetry) follows from the fact  $\sum_{i'' \in N(j) \setminus \{i\}} \gamma_{i''} = \sum_{i'' \in N(j) \setminus \{i'\}} \gamma_{i''}$  for  $\gamma_i = \gamma_{i'} = 1$ . The first property (marginalization) can be verified directly.  $\Box$ 

Remark 2. The function  $\rho_j$  has a natural probabilistic interpretation. As in the proof, we can think of z as the mean of a distribution over assignments  $\gamma \in C_j$ . For a variable node  $i \in N(j)$  with  $z_i > 0$ , we sample an assignment  $\gamma$  from this distribution conditioned on the event  $\gamma_i = 1$ . Now we output a random variable node  $i' \in N(j) \setminus \{i\}$  with  $\gamma_{i'} = 1$ . The probability that we output i' is exactly  $\rho_j(i, i')/z_i$ . Note that the function  $\rho_j$  is not fully determined by z, since we could realize z as mean of very different distributions.

The distribution over minimal *T*-local deviations  $\beta$  we construct for the LP solution *z* is very similar to the distribution used in Lemma 1 (especially taking into account the probabilistic interpretation of the functions  $\rho_j$ ). As before, we first define for every variable node *i* with  $z_i > 0$ , a distribution over height-2*T* skinny trees  $\tau_i$  rooted at *i*: we start by choosing *i* as the root. From any chosen variable node, we branch to all its neighbors in the next level. From any chosen check node *j*, we go to a random neighbor, chosen according to the transition probabilities  $\rho_j(\cdot, \cdot)$ . More precisely, if we reached *j* from a variable node  $i_{\rm in}$ , then we select the next variable node  $i_{\rm out}$  at random from  $N(j) \setminus \{i_{\rm in}\}$ , with probability  $\rho_j(i_{\rm in}, i_{\rm out})/z_{i_{\rm in}}$ .

We can now define the distribution over minimal local deviations  $\beta$ : For every variable node *i* with  $z_i > 0$ , we independently choose a tree  $\tau_i$  as described above. Independently from the choices of the trees, we also choose a variable node  $i_0$  at random according to the probabilities  $z_{i_0}/||z||_1$ . Finally, we output the minimal *T*-local deviation  $\beta$  defined by

$$\beta_i = \begin{cases} 1 & \text{if } i \in \tau_{i_0} ,\\ 0 & \text{otherwise} . \end{cases}$$

We make a few observations. First, the number of nodes at level 2t of  $\tau_{i_0}$  is exactly  $d_L(d_L - 1)^{t-1}$ . Second, for any two variable nodes i, i' that lie in the support of z and have distance  $\leq 2T$  to each other, the above process for constructing a random skinny tree treats them symmetrically:

$$z_i \mathbb{P}\left\{i' \in \tau_i\right\} = z_{i'} \mathbb{P}\left\{i \in \tau_{i'}\right\} \tag{7}$$

The reason is that i, i' are connected by a single path of length  $\leq 2T$ , say  $(i, j_0, i_1, j_1, \ldots, i_{t-1}, j_{t-1}, i')$ . If  $z_{i_{\ell}} = 0$  for some variable node  $i_{\ell}$  on this path, both sides of (7) are naught. Otherwise, both sides of (7) are equal to

$$rac{
ho_{j_0}(i,i_1)\cdots
ho_{j_{t-1}}(i,i')}{z_{i_1}\cdots z_{i_{t-1}}}$$
 .

Armed with these observations, we can compute the mean of our distribution over (w-weighted) minimal local deviations: For every i with  $z_i > 0$ , we have

$$\mathbb{E}\beta_{i}^{(w)} = \sum_{t=1}^{T} w_{t} \sum_{\substack{i' \in N^{2t}(i) \\ z_{i'} > 0}} \mathbb{P}\left\{i_{0} = i'\right\} \mathbb{P}\left\{i \in \tau_{i'}\right\}$$
$$= \sum_{t=1}^{T} w_{t} \sum_{\substack{i' \in N^{2t}(i) \\ z_{i'} > 0}} \frac{z_{i'}}{\|z\|_{1}} \mathbb{P}\left\{i \in \tau_{i'}\right\}$$
$$\stackrel{(7)}{=} \sum_{t=1}^{T} w_{t} \sum_{\substack{i' \in N^{2t}(i) \\ z_{i'} > 0}} \frac{z_{i}}{\|z\|_{1}} \mathbb{P}\left\{i' \in \tau_{i}\right\}$$
$$= z_{i} \cdot \frac{1}{\|z\|_{1}} \sum_{t=1}^{T} w_{t} \mathbb{E}\left|\tau_{i} \cap N^{2t}(i)\right|$$
$$= z_{i} \cdot \frac{1}{\|z\|_{1}} \sum_{t=1}^{T} w_{t} \cdot d_{L}(d_{L} - 1)^{t-1}$$

Therefore, we have the desired conclusion  $\mathbb{E}\beta^{(w)} = \alpha z$ with  $\alpha = \sum_{t=1}^{T} w_t \cdot d_L (d_L - 1)^{t-1} / ||z||_1$ .  $\Box$ 

# 6. PROBABILISTIC ANALYSIS OF LO-CAL OPTIMALITY ON TREES

For the purposes of this section, let us define a notion of optimality of a codeword in the immediate neighborhood of a variable node  $i_0 \in V_L$ , by appropriately restricting Definition 2 of Section 3.

Definition 3. (Single Neighborhood Optimality) A codeword  $x \in \{0, 1\}^n$  is  $(i_0, T, w)$ -locally optimal for  $y \in \{0, 1\}^n$ if for all minimal T-local deviations  $\beta$  at  $i_0$ ,

$$||x \oplus \beta^{(w)} - y||_1 > ||x - y||_1$$

Now, for a fixed weight vector w, variable node  $i_0 \in V_L$ , and codeword  $x \in C(G)$ , we are interested in the probability

$$\mathbb{P}_{y \sim_p x} \left\{ \begin{array}{c} x \text{ is } (i_0, T, w) \text{-locally} \\ \text{optimal for } y \end{array} \right\}, \tag{8}$$

where  $\mathbb{P}_{y \sim_p x}$  is the measure defined by flipping every bit of x independently with probability p to obtain y.

We show that, if  $T < \frac{1}{4} \operatorname{girth}(G)$ , the symmetry of the code and the channel imply that the probability in (8) does not depend on x and  $i_0$ . Therefore, estimating this probability becomes a very concrete question about a random process in a fixed regular tree (Definition 4 in box).

The following lemma makes the connection of this random process to local optimality precise.

$$\mathbb{P}_{y \sim_p x} \left\{ \begin{array}{c} x \ is \ (i_0, T, w) \text{-locally} \\ optimal \ for \ y \end{array} \right\} = \mathbb{P}_p \left\{ \min_{\tau} \operatorname{val}_{\omega}(\tau; \eta) > 0 \right\},$$

where  $\omega_{\ell} = w_{T-\ell}$ .

LEMMA 5. Let  $T < \frac{1}{4} \operatorname{girth}(G)$ .

PROOF. The subgraph of G in  $B(i_0, 2T)$  is isomorphic to the tree  $\mathcal{T}$ . Let  $\varphi: B(i_0, 2T) \to V(\mathcal{T})$  be one of the isomorphisms between the two graphs.

First, we observe that x is  $(i_0, T, w)$ -locally optimal for y if and only if

$$\min_{\beta} \sum_{t=1}^{T} \sum_{i \in N^{2t}(i_0)} w_t \cdot (-1)^{x_i + y_i} \beta_i > 0,$$

Definition 4.  $(T, \omega)$ -Process on a  $(d_L, d_R)$ -Tree: Let  $\mathcal{T}$  be a directed tree of height 2T, rooted at a vertex  $v_0$ . The root has out-degree  $d_L$  and the vertices in level 2T have out-degree 0. The vertices in any other even level have out-degree  $d_L - 1$ . The vertices in odd levels have out-degree  $d_R - 1$ . The vertices in even levels are called variable nodes and the vertices in odd levels are called check nodes. For  $\ell \in \{0, \ldots, 2T\}$ , let us denote by  $V_\ell$  the set of vertices of  $\mathcal{T}$  at height  $\ell$  (the leaves have height 0 and the root has height 2T).

A skinny subtree of  $\mathcal{T}$  is a vertex set  $\tau \subseteq V(\mathcal{T})$  such that the induced subgraph is a connected tree containing the root  $v_0$  where each variable node in  $\tau$  has full out-degree and each check node in  $\tau$  has out-degree exactly 1. For a  $\{1, -1\}$ -assignment  $\eta$  to the variable nodes of  $\mathcal{T}$ , we define the  $\omega$ -weighted value of a skinny subtree  $\tau$  as

$$\operatorname{val}_{\omega}(\tau;\eta) \stackrel{\mathrm{def}}{=} \sum_{\ell=0}^{T-1} \sum_{v \in \tau \cap V_{2\ell}} \omega_{\ell} \cdot \eta_{v}$$

(In words, we sum the values of the variable nodes in  $\tau$  weighted according to their height.)

For  $p \in (0, 1)$ , we are interested in the probability

$$\mathbb{P}_p\left\{\min_{\tau} \operatorname{val}_{\omega}(\tau;\eta) > 0\right\} \,,$$

where the minimum is over all skinny subtrees  $\tau$  and the measure  $\mathbb{P}_p$  on  $\eta$  is defined by choosing  $\eta_v = 1$  with probability 1 - p, and  $\eta_v = -1$  with probability p.

where the minimum is over all minimal *T*-local deviations  $\beta$  at  $i_0$ . The reason is that  $\|x \oplus \beta^{(w)} - y\|_1 - \|x - y\|_1$  can be expanded as  $\sum_{i=1}^n \beta_i^{(w)} (-1)^{x_i+y_i}$ .

We also note that the isomorphism  $\varphi$  gives rise to a bijection between the minimal deviations  $\beta$  in  $B(i_0, 2T)$  and the skinny subtrees  $\tau$  of  $\mathcal{T}$ . We define  $\varphi(\beta)$  to be the skinny tree that contains all variable nodes v with  $\beta_{\varphi^{-1}(v)} = 1$ .

We can finish the proof by coupling the random variables y and  $\eta$  in such a way that  $||x \oplus \beta^{(w)} - y||_1 - ||x - y||_1 =$ val $_{\omega}(\varphi(\beta); \eta)$ . We use the following coupling

$$\eta_v = (-1)^{x_i + y_i}$$
, where  $i = \varphi^{-1}(v)$ .

We verify that for all  $\beta$  and  $\tau = \varphi(\beta)$ 

$$\begin{aligned} \|x \oplus \beta^{(w)} - y\|_{1} - \|x - y\|_{1} \\ &= \sum_{t=1}^{T} \sum_{i \in N^{2t}(i_{0})} w_{t} \cdot (-1)^{x_{i} + y_{i}} \beta_{i} = \sum_{t=1}^{T} \sum_{i \in N^{2t}(i_{0})} w_{t} \cdot \eta_{\varphi(i)} \beta_{i} \\ &= \sum_{t=1}^{T} \sum_{v \in \tau \cap V_{2T-2t}} w_{t} \cdot \eta_{v} = \sum_{\ell=0}^{T-1} \sum_{v \in \tau \cap V_{2\ell}} \omega_{\ell} \cdot \eta_{v} \\ &= \operatorname{val}_{\omega}(\phi(\beta); \eta) \quad \Box \end{aligned}$$

Let us define

$$\Pi_{p,d_L,d_R}(T,\omega) \stackrel{\text{def}}{=} \mathbb{P}_p\left\{\min_{\tau} \operatorname{val}_{\omega}(\tau;\eta) \leqslant 0\right\}$$

With this notation, Lemma 5 together with Theorem 4 (Local optimality implies LP optimality) has the following consequence.

LEMMA 6. Let  $p \in (0, 1)$ , G be a  $(d_L, d_R)$ -regular bipartite graph,  $x \in C(G)$  be a codeword, and  $w \in [0, 1]^T$  be a weight vector with  $T < \frac{1}{4} \operatorname{girth}(G)$ . Suppose y be obtained from x by flipping every bit independently with probability p. Then, codeword x is (T, w)-locally optimal with probability at least

$$1 - n \cdot \prod_{p,d_L,d_R} (T,\omega)$$
, where  $\omega_\ell = w_{T-\ell}$ .

And with at least the same probability, x is also the unique optimal LP solution given y.

By virtue of Lemma 6, to understand the probability of LP decoding success, it is sufficient to estimate the probability  $\Pi_{p,d_L,d_R}(T,\omega)$ , for a given weight vector w, bit error rate  $p \in (0,1)$ , and degrees  $(d_L, d_R)$ . We give such estimates in the following subsection.

## 6.1 Bounding Processes on Trees by Evolving Laplace Transforms

We are going to study the probability of the existence of a negative value skinny subgraph in the  $(T, \omega)$ -process in a recursive fashion, starting from the leaves of the tree  $\mathcal{T}$ .

We define the following correlated random variables  $Z_u$  for the vertices u of  $\mathcal{T}$ : The variable  $Z_u$  is equal to the minimum value of a skinny tree in the subtree  $\mathcal{T}_u$  below the vertex u,

$$Z_u \stackrel{\text{def}}{=} \min_{\tau} \sum_{\ell=0}^{T-1} \sum_{v \in \tau \cap V_{2\ell} \cap \mathcal{T}_u} \omega_\ell \cdot \eta_v \,.$$

Here,  $\tau$  ranges over all skinny subtrees of  $\mathcal{T}$ .

Let  $N^+(u)$  denote the set of neighbors of u that can be reached by one of its outgoing edges. The variables  $Z_u$  satisfy the following recurrence relations:

$$Z_{v_0} = \sum_{v \in N^+(v_0)} Z_v$$

$$Z_u = \omega_\ell \eta_u + \sum_{v \in N^+(u)} Z_v \quad (u \in V_{2\ell}, \ 0 \le \ell < T)$$

$$Z_u = \min_{v \in N^+(u)} Z_v \quad (u \in V_{2\ell+1}, \ 0 \le \ell < T)$$

Note that  $Z_{v_0}$  is just the minimum value of a skinny tree in the tree  $\mathcal{T}$ . Hence,  $\prod_{p,d_L,d_R}(T,\omega) = \mathbb{P}\{Z_{v_0} \leq 0\}.$ 

By symmetry, the distribution of a variable  $Z_u$  depends only on the height of vertex u. Also the variables in  $\{Z_u\}_{u \in V_\ell}$ are mutually independent, because for any two vertices u, u'of the same height  $\ell$ , the subtrees  $\mathcal{T}_u$  and  $\mathcal{T}_{u'}$  are disjoint.

It follows that we can define (uncorrelated) random variables  $X_0, \ldots, X_{T-1}, Y_0, \ldots, Y_{T-1}$  in the following way, so that  $X_{\ell}$  has the same distribution as  $Z_u$  for  $u \in V_{2\ell+1}$  and  $Y_{\ell}$  has the same distribution as  $Z_u$  for  $u \in V_{2\ell}$ ,

$$Y_{0} = \omega_{0}\eta$$

$$X_{\ell} = \min\left\{Y_{\ell}^{(1)}, \dots, Y_{\ell}^{(d_{R}-1)}\right\} \quad (0 \leq \ell < T)$$

$$Y_{\ell} = \omega_{\ell}\eta + X_{\ell-1}^{(1)} + \dots + X_{\ell-1}^{(d_{L}-1)} \quad (0 < \ell < T)$$

Here,  $\eta$  is a random variable that takes value 1 with probability 1-p and value -1 with probability p. The notation  $X^{(1)}, \ldots, X^{(d)}$  means that we take d mutually independent copies of the random variable X (the copies are also independent of  $\eta$ ).

We will use the Laplace transform of  $X_{T-1}$  in order to bound the probability  $\prod_{p,d_L,d_R}(T,\omega)$ .

Lemma 7. For every 
$$t \ge 0$$
,  

$$\Pi_{p,d_L,d_R}(T,\omega) \leqslant \left(\mathbb{E}\,e^{-tX_{T-1}}\right)^{d_L} \,.$$

PROOF. As noted before,  $\prod_{p,d_L,d_R}(T,\omega) = \mathbb{P}\{Z_{v_0} \leq 0\}$ . Hence, by Markov's inequality

$$\Pi_{p,d_L,d_R}(T,\omega) = \mathbb{P}\left\{e^{-tZ_{v_0}} \geqslant 1\right\} \leqslant \mathbb{E} e^{-tZ_{v_0}} .$$

The variable  $Z_{v_0}$  is equal to the sum of the Z-values of its  $d_L$  children. Each child of the root  $v_0$  has height 2T - 1 and hence its Z-value has the same distribution as  $X_{T-1}$ . Thus, we have as desired

$$\mathbb{E} e^{-tZ_{v_0}} = \left(\mathbb{E} e^{-tX_{T-1}}\right)^{d_L} . \quad \Box$$

The following is our key lemma for estimating the probability  $\prod_{p,d_L,d_R}(T,\omega)$  (or more precisely, the Laplace transform of  $X_{\ell}$ ). For the sake of brevity, let us denote  $d'_L = d_L - 1$ and  $d'_R = d_R - 1$ .

LEMMA 8. For  $\ell, s$  with  $0 \leq s \leq \ell < T$ , we have

$$\mathbb{E} e^{-tX_{\ell}} \leqslant \left(\mathbb{E} e^{-tX_s}\right)^{d'_L^{\ell-s}} \cdot \prod_{k=0}^{\ell-s-1} (d'_R \mathbb{E} e^{-t\omega_{\ell-k}\eta})^{d'_L^k}.$$

PROOF. We derive the relation for  $s = \ell - 1$ . The general case follows by induction on  $\ell - s$ .

Since  $Y_{\ell}$  is a sum of mutually independent variables,

$$\mathbb{E} e^{-tY_{\ell}} = \left(\mathbb{E} e^{-t\omega_t \eta}\right) \left(\mathbb{E} e^{-tX_{\ell-1}}\right)^{d'_L}$$

We use a relatively crude estimate to bound the Laplace transform of  $X_{\ell}$  in terms of the Laplace transform of  $Y_{\ell}$ . By the definition of  $X_{\ell}$ , we have  $\exp(-tX_{\ell}) \leq \exp(-tY_{\ell}^{(1)}) + \dots + \exp(-tY_{\ell}^{(d_R-1)})$  with probability 1. Hence,

$$\mathbb{E} e^{-tX_{\ell}} \leqslant d'_R \mathbb{E} e^{-tY_{\ell}} = \left(d'_R \mathbb{E} e^{-t\omega_{\ell}\eta}\right) \left(\mathbb{E} e^{-tX_{\ell-1}}\right)^{d'_L},$$

which is the desired bound for  $s = \ell - 1$ .  $\Box$ 

Armed with these general bounds on  $\prod_{p,d_L,d_R}(T,\omega)$  and the Laplace transform of  $X_{\ell}$ , we can now derive several concrete bounds on  $\prod_{p,d_L,d_R}(T,\omega)$ .

#### Uniform Weights.

In this paragraph, we will consider the case  $\omega = 1$ . We apply Lemma 8 for s = 0. For brevity, let us denote  $c_1 = \mathbb{E} e^{-tX_0}$  and  $c_2 = d'_R \mathbb{E} e^{-t\eta}$ . Note that  $c_1 \leq c_2$  (same argument as in the proof of Lemma 8). For reasons that become apparent shortly, let us choose  $t \geq 0$  so as to minimize  $c := c_1 \cdot c_2^{1/(d_L-2)}$ . We will assume c < 1. Now, the bound of Lemma 8 simplifies to

$$\mathbb{E} e^{-tX_{\ell}} \leqslant c_1^{d'_{L}^{\ell}} \cdot c_2^{\sum_{k=0}^{\ell-1} d'_{L}^{k}} = c_1^{d'_{L}^{\ell}} \cdot \left(c_2^{1/(d'_{L}-1)}\right)^{d'_{L}^{\ell}-1}$$
$$= c^{d'_{L}^{\ell}} \cdot c_2^{-1/(d'_{L}-1)} \leqslant c^{d'_{L}^{\ell}-1}.$$

By Lemma 7 we can conclude from this bound that

$$\Pi_{p,d_L,d_R}(T,\mathbb{1}) \leqslant c^{d_L d'_L^{-1} - d_L}$$

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Next, let us compute  $c_1$  and  $c_2$  as functions of p,  $d_L$  and  $d_R$ . The variable  $X_0$  has the following distribution

$$X_0 = \begin{cases} +1, & \text{with probability } (1-p)^{d_R-1}, \\ -1, & \text{with probability } 1-(1-p)^{d_R-1}, \end{cases}$$

Hence,

$$c_1 = \mathbb{E} e^{-tX_0} = (1-p)^{d_R-1} e^{-t} + (1-(1-p)^{d_R-1}) e^t.$$

We also have

$$c_2 = (d_R - 1) \left( (1 - p)e^{-t} + pe^t \right) .$$

Putting together the calculations in this paragraph, we proved the following general bound on  $\Pi_{p,d_L,d_R}(T, 1)$ .

LEMMA 9. If  $p \in (0,1)$  and  $d_L, d_R \ge 2$  satisfy the condition

$$c = \min_{t \ge 0} \left( (1-p)^{d_R-1} e^{-t} + (1-(1-p)^{d_R-1}) e^t \right) \\ \cdot \left( (d_R-1) \left( (1-p)e^{-t} + pe^t \right) \right)^{1/(d_L-2)} < 1,$$

then for  $T \in \mathbb{N}$  and  $w = (1, \dots, 1) \in [0, 1]^T$ , we have

$$\prod_{p,d_L,d_R} (T,\omega) \leqslant c^{d_L d'_L T - 1} - d_L \,.$$

For (3, 6)-regular graphs, we have the following corollary.

COROLLARY 1. Let  $p \leq 0.02$ ,  $d_L = 3$ , and  $d_R = 6$ . Then, there exists a constant c < 1 such that for all T and  $\omega = 1$ ,

$$\Pi_{p,d_L,d_R}(T,\omega) \leqslant c^{2^1}$$

Non-uniform Weights.

In this paragraph, we will show how to improve the bounds by using different weights according to the height. We will use very simple weights: variable nodes at height 0 are weighted by a factor  $\omega_0 \ge 0$ , all variable nodes at higher heights are weighted by 1.

We apply Lemma 8 again for s = 0. As in the previous paragraph, the bound simplifies to

$$\mathbb{E} e^{-tX_{\ell}} \leqslant c^{d'_L^{\ell}} \cdot c_2^{-1/(d'_L - 1)}$$

where  $c_1 = \mathbb{E} e^{-tX_0}$ ,  $c_2 = d'_R \mathbb{E} e^{-t\eta}$ , and  $c = c_1 \cdot c_2^{1/(d_L-2)}$ . The additional freedom of choosing the weight  $\omega_0$ , allows us to minimize both  $c_1$  and  $c_2$  at the same time. To minimize  $c_2$ , we need to choose  $t = \frac{1}{2} \ln \frac{1-p}{p}$ . The value of  $c_1$  is equal to

$$c_1 = \mathbb{E} e^{-tX_0} = (1-p)^{d_R-1} e^{-t\omega_0} + (1-(1-p)^{d_R-1}) e^{t\omega_0},$$

which is minimized for

$$t\omega_0 = \ln \sqrt{\frac{(1-p)^{d_R-1}}{1-(1-p)^{d_R-1}}}$$

Here, the right-hand side is nonnegative for  $p < 1 - 2^{-1/d'_R}$ . (Note that we do not have to worry whether  $\omega_0 \leq 1$ . By the definition of the  $(T, \omega)$ -process,  $\prod_{p,d_L,d_R}(T, \omega)$  is invariant under (nonnegative) scaling of the weights.)

For these choices of  $\omega_0$  and t, we have

$$c_1 = 2\sqrt{(1-p)^{d_R-1} (1-(1-p)^{d_R-1})}$$
  
$$c_2 = d'_R 2\sqrt{p(1-p)}.$$

Thus,

$$c = 2\sqrt{(1-p)^{d'_R} \left(1-(1-p)^{d'_R}\right)} \left(d'_R 2\sqrt{p(1-p)}\right)^{-1/(d_L-2)}.$$

We proved the following bound on  $\Pi_{p,d_L,d_R}(T,\omega)$  for  $\omega = (\omega_0, 1, \ldots, 1)$ .

LEMMA 10. If  $p \in (0,1)$  and  $d_L, d_R \ge 2$  satisfy the condition  $p < 1 - 2^{-1/d'_R}$  and  $c_{p,d_L,d_R} < 1$ , where

$$c_{p,d_L,d_R} \stackrel{\text{def}}{=} 2\sqrt{(1-p)^{d'_R} \left(1-(1-p)^{d'_R}\right)} \left(d'_R 2\sqrt{p(1-p)}\right)^{1/(d_L-2)}$$

then there exists a constant  $\omega_0 \ge 0$  such that for all  $T \in \mathbb{N}$ and  $\omega = (\omega_0, 1, \dots, 1)$ ,

$$\Pi_{p,d_L,d_R}(T,\omega) \leqslant c' c^{d_L d'_L^{T-1}},$$

where 
$$c' = (d'_R 2 \sqrt{p(1-p)})^{-(d_L-1)/(d_L-2)}$$
.

COROLLARY 2. Let  $p \leq 0.0247$ ,  $d_L = 3$ , and  $d_R = 6$ . Then, there exists a constant c < 1 such that for all T,

$$\Pi_{p,d_L,d_R}(T,\omega) \leqslant c^{2^T} \quad \text{for some } \omega \in [0,1]^T.$$

#### *Improved Bounds for* (3,6)-*Regular Trees.*

In this paragraph we show how to obtain the bound of 0.05 on the tolerable noise rate of (3, 6)-regular codes.

Computer simulations suggest that for p > 0.025, the probability  $\prod_{p,3,6}(T, 1)$  approaches 1 as  $T \to \infty$ . Thus, it seems necessary to use non-uniform weights in order to achieve the bound 0.05.

Let us first consider the weight vector  $\bar{\omega} = (1, 2, \dots, 2^s)$ . Note that this weight vector has the effect that every level contributes equally to the  $\omega$ -weighted value  $\operatorname{val}_{\bar{\omega}}(\tau;\eta)$  of a skinny subtree  $\tau$ . For a concrete value of s (say s = 15), we can compute the distribution of  $X_s$  using the recursive definition of the X and Y variables. Hence, for a fixed s, we can also compute the value

$$\lambda_s \stackrel{\text{def}}{=} \min_{t \ge 0} \mathbb{E} e^{-tX_s}$$

Let  $t^* \ge 0$  be the point where the Laplace transform of  $X_s$  achieves its minimum  $\lambda_s$ . We now show how to bound  $\prod_{p,3,6}(T,\omega)$  in terms of  $\lambda_s$  for  $\omega = (\bar{\omega}, \rho, \ldots, \rho) \in \mathbb{R}^T_+$ , where  $\rho$  is a carefully chosen constant.

By Lemma 8, we have

$$\mathbb{E} e^{-t^* X_{T-1}} = (\lambda_s)^{2^{T-s-1}} \left( 5 \mathbb{E} e^{-t^* \rho \eta} \right)^{\sum_{k=0}^{T-s-2} 2^k} \\ = (\lambda_s)^{2^{T-s-1}} \left( 5 \mathbb{E} e^{-t^* \rho \eta} \right)^{2^{T-s-1}-1} \\ = \left( \lambda_s \cdot 10 \sqrt{p(1-p)} \right)^{2^{T-s-1}} \left( 10 \sqrt{p(1-p)} \right)^{-1} ,$$

where we chose  $\rho$  such that  $e^{t^*\rho} = \sqrt{(1-p)/p}$  (then  $\mathbb{E} e^{-t^*\rho\eta}$  is minimized). Using 7, we can see that  $\Pi_{p,3,6}(T,\omega)$  decrease doubly-exponential in T if  $\lambda_s \cdot 10\sqrt{p(1-p)} < 1$  for some s. We verified that this condition is satisfied for s = 15 and p = 0.05 using the numerical analysis software MATLAB.

We extend this approach to general  $(d_L, d_R)$ -regular trees in the following lemma.

LEMMA 11. Let  $p \in (0, \frac{1}{2})$  and  $d_L, d_R \ge 2$ . Suppose that for some  $s \in \mathbb{N}$  and some weight vector  $\bar{\omega} \in \mathbb{R}^{s}_{+}$ ,

$$\min_{t \ge 0} \mathbb{E} e^{-tX_s} < \frac{1}{(d_R - 1)2\sqrt{p(1-p)}}$$

Then, there exists constants c < 1 and  $\rho \ge 0$  such that for all T,

$$\Pi_{p,d_L,d_R}(T,\omega) \leqslant c^{d_L(d_L-1)^{T-1}},$$
  
where  $\omega = (\bar{\omega}, \rho, \dots, \rho) \in \mathbb{R}_+^T.$ 

COROLLARY 3. Let  $p \leq 0.05$ ,  $d_L = 3$ , and  $d_R = 6$ . Then, there exists a constant c < 1 such that for all T,

$$\Pi_{p,d_L,d_R}(T,\omega) \leqslant c^{2^T} \quad \text{for some } \omega \in [0,1]^T \,.$$

# 7. CONCLUSIONS

Our original intention was to connect Belief Propagation to Linear Programming (or some other form of convex programming) and this remains open. It is unclear where to start since BP relies on highly nonlinear operations.

It would also be interesting to investigate if stronger versions of LP decoding using either lift-and-project operators such as Sherali Adams or using SDPs could have better provable performance for LDPCs, possibly approaching the information theoretic bound.

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